

SCHOOL OF OPERATIONS RESEARCH
AND INDUSTRIAL ENGINEERING
COLLEGE OF ENGINEERING
CORNELL UNIVERSITY
ITHACA, NY 14853-3801

TECHNICAL REPORT NO. 1095

May 1994
Revised June 1994

**A COMPUTER PROGRAM
FOR THE STATISTICAL ANALYSIS
OF REPEATED EVENT DATA
USING A MIXED EFFECTS
REGRESSION MODEL¹**

by

R. Natarajan, B.W. Turnbull, E.H. Slate,
M.T. Wells, L.C. Clark, and H. Abu-Libdeh

¹This research was supported principally by Grant R01 CA61120, also by grants R01 GM28364 and R01 CA 49764, all from the U.S. National Institutes of Health.

A COMPUTER PROGRAM FOR THE STATISTICAL ANALYSIS OF REPEATED EVENT DATA USING A MIXED EFFECTS REGRESSION MODEL

Ranjini NATARAJAN, Bruce W. TURNBULL, Elizabeth H. SLATE, Martin T. WELLS,

Cornell University, Ithaca NY 14853

Larry C. CLARK

University of Arizona

and

Hasan ABU-LIBDEH

Department of Mathematics, Birzeit University, West Bank, via Israel

This paper presents a computer program for fitting mixed effects regression models to repeated events data. The method has been described by Abu-Libdeh, Turnbull and Clark [1]. Such data can occur in longitudinal studies where subjects experience repeated events over time. The program allows the stepwise construction of a series of regression models which can be used to examine and test the influence of the various measured covariates upon the event rates. Two examples are provided. The first is a simple example involving the incidence of mammary tumors in rats. The second involves a very large complex data set from a clinical trial for the prevention of recurrent skin tumors.

Key Words: Multiple failure times, Covariates, Poisson regression, Random effects.

Corresponding Author: Bruce W. Turnbull,

227 ETC, School of Operations Research and Industrial Engineering,
Cornell University, Ithaca, New York 14850-3801, USA.

E-mail: turnbull@orie.cornell.edu

June 9, 1994

1 Introduction

The purpose of this paper is to present and describe a computer program which performs a stepwise (forward and/or backward) procedure for fitting regression models to repeated events data. Such data can occur in longitudinal studies, either experimental or observational, where individuals may experience repeated events (e.g. failures) over time. The events can be represented as a realization of a stochastic point process and some researchers have alternately described this situation as one with multiple failure time data. Individuals are assumed heterogeneous. Some of the heterogeneity is explained by measured covariates (e.g. treatment, age, gender); unexplained subject variability is modeled by using random effects. There are applications to many fields in biomedicine: in clinical studies, for example, of epileptic seizures, asthma, skin cancer (Abu-Libdeh, Turnbull and Clark [1]), gallstones (Thall [2]), bladder cancer (Freedman, Sylvester and Byar [3]); in animal carcinogenicity experiments to incidence of mammary tumors (Gail, Santner and Brown[4]). There are also applications in engineering to the reliability of repairable systems (Ascher and Feingold [5]) and in sociology to studies of mobility, unemployment, fertility, etc. (Allison [6]).

If at most one event (failure, death etc.) can occur to any one individual, then the survival analysis techniques such as exponential, Weibull or proportional hazards regression methods are applicable (see, e.g. Lawless 1982). For the situation where repeated events can occur over time in a single individual, as in some of the studies listed above, several statistical methods have been proposed by various researchers. This paper concerns a quite flexible parametric approach, based on a Poisson process mixture model, described by Lawless [8, Sec. 3] and later expanded by Abu-Libdeh et al. [3]. The latter applied it to data from an ongoing randomized multi-center clinical trial for the prevention of reoccurrence of skin cancer.

The methodology is described in Section 2. The computer program is described in Section 3. Some sample runs are described in Section 4. Hardware and software specifications are given in Section 5. The program is written in FORTRAN and versions are currently set up to run interactively on a SUN 4 workstation and on an IBM PC or compatible.

2 Description of the Methods

We suppose there are a total of I individuals or subjects. For individual i , $1 \leq i \leq I$, there are n_i events, possibly zero, over a period of observation T_i . The period of observation is the elapsed time from that individual's entry time on the study to the time last seen. These $\{T_i\}$ may vary from individual to individual. A set of p (≥ 1) "regressor", "explanatory" or "covariate" values, denoted by the row p -vector \mathbf{x}_i , is associated with each subject i , $1 \leq i \leq I$. These are the so-called "fixed effects" and might represent such variables as age, gender, treatment, smoking status etc.

The model specifies that the number of events n_i experienced by individual i , $1 \leq i \leq I$, is Poisson distributed with mean $\mu_i = \theta_i T_i \exp(\mathbf{x}_i \boldsymbol{\beta})$, i.e. proportional to the follow-up time T_i . Here $\boldsymbol{\beta}$ is a (column) p -vector of regression coefficients, representing the effect of the regressor variables, and θ_i is the baseline ($\mathbf{x}_i = \mathbf{0}$) value or "underlying" event rate for the i th subject. Equivalently, we may say that the events for individual i occur according as a Poisson process with event rate $\theta_i \exp(\mathbf{x}_i \boldsymbol{\beta})$. Another characterization is that the inter-event times are independent exponentially distributed random variables with means given by $[\theta_i \exp(\mathbf{x}_i \boldsymbol{\beta})]^{-1}$.

Subject heterogeneity (or "extra Poisson variation") that is not explained by the regressor variables is modeled by allowing the θ_i values to vary from subject to subject according to a gamma distribution with scale parameter γ and shape parameter ν (Johnson and Kotz, [9, p. 166, eqn (2)]). Thus unexplained patient variability is modeled as a random effect. The gamma mixing distribution is a convenient and flexible choice that has been used by Lawless [8] and others – see the remark by Abu-Libdeh et al. [3, p.1019]. The mean and variance of the mixing distribution are given by $\mu = \nu\gamma$ and $\nu\gamma^2$, respectively. Of special interest is the degenerate case when $\nu \rightarrow \infty$ so that the θ_i are equal and constant and the baseline subject event rates can be considered homogeneous. We have constructed this model as a mixed (random and fixed) effects model; however it may be viewed in a number of ways, such as a negative binomial regression model, as a generalized linear model, or as a model with "extra-Poisson" variation (Lawless [8,10]).

Conditional on θ_i , the contribution to the likelihood for the i th subject is

$$L_i(\boldsymbol{\beta} | \theta_i) = \mu_i^{n_i} e^{-\mu_i} / n_i!$$

where $\mu_i = \theta_i T_i \exp(\mathbf{x}_i \boldsymbol{\beta})$,

The full likelihood is then given by

$$L(\gamma, \nu, \beta) = \prod_{i=1}^I \int_0^\infty L_i(\beta | \theta_i) g(\theta_i; \gamma, \nu) d\theta_i$$

where $g(\theta; \gamma, \nu) = \gamma^{-\nu} \theta^{\nu-1} \exp(-\theta/\gamma)/\Gamma(\nu)$ is the gamma density of the mixing distribution.

Carrying out the integration we obtain;

$$L(\gamma, \nu, \beta) = \prod_{i=1}^I \frac{\Gamma(n_i + \nu)}{\Gamma(\nu)} \frac{[\gamma \exp(\mathbf{x}_i \beta)]^{n_i}}{[\gamma T_i \exp(\mathbf{x}_i \beta) + 1]^{n_i + \nu}} \quad (1)$$

This likelihood, in a slightly different form, is the same as that given by Lawless (1987b, p.210). It is also a special case of Abu-Libdeh et al. (1990, Eqn (1)). For numerical reasons, it is easier to reparametrize and work with parameters $\mu = \nu\gamma$ and ν rather than γ and ν . Replacing γ by μ/ν in (1), we write the resulting likelihood as $L(\mu, \nu, \beta)$

Estimates of the mixing distribution parameters μ, ν (equivalently γ, ν) and of the regression coefficients $\beta = (\beta_1, \dots, \beta_p)$ can be found by maximizing the likelihood (1), or equivalently its logarithm $\log L = \mathcal{L}$, say, with respect to these variables. The maximizing values are the maximum likelihood estimates. The variance-covariance matrix of the estimates is obtained by evaluating the inverse of the information matrix at the estimated parameter values. The information matrix is given by the negative of the matrix of second derivatives (Hessian) of $\mathcal{L}(\mu, \nu, \beta)$. The approximate standard errors of the estimates are then given by the square roots of the diagonal entries of the variance-covariance matrix. This leads to construction of confidence intervals for each of the parameters estimated and to the Wald statistic for testing the significance of each of the estimates. Alternatively, a score statistic, based on the gradient of \mathcal{L} , or a likelihood ratio statistic can be used to test the significance of a regression coefficient – see Rao [11, Sec. 6e.3] or Cox and Hinkley [12].

3 Program Description

A flow chart for the main program is given in Figure 1. Some sample runs are described in the next section. The program consists of four parts (1) Data input; (2) Model specification; (3) Likelihood maximization; and (4) Output of parameter estimates, standard errors and test statistics. The user has the option then to return to step (2) to specify an enlarged set of regressor variables to

include in the model. Thus a preferred model can be constructed in a forward stepwise manner by repeating steps (2)–(4).

3.1 Step 1: Data input

The user is first prompted for the name DICTION of the “dictionary” file. The first line of this file must contain the name of the DATA file. The second line contains the values of NVARS+1 (ID and the number of regressor variables), and NSUBS, the number of subjects (I). Lines 3 through NVARS+3 contain the descriptive names of the record identifier (ID), the number of events, the observation time, and each of the regressor variables (e.g. “age”, “treatment”, “smoking status” etc. in the example of Section 4.2). The dictionary file for the small example of Section 4.1 is shown in Table 1.

The DATA file named in the first line of the DICTION file is now read in. This is a rectangular file with NSUBS rows and NVARS+3 columns in free format. In the i th row, which corresponds to subject i , the first entry is its name or “ID”. This item is not actually used by the program, but its presence is useful to detect errors in data entry. The second entry in the row is the number of events n_i experienced by the i th subject; the third item is the observation, followup or exposure time T_i for that subject. Entries 4 through NVARS+3 are the covariate values for subject i in the same order as that specified in the dictionary file. Missing values are designated by the code “-999”. Missing values in a regressor variable cause the corresponding records to be ignored for models including this variable. Missing (-999) or negative values for the number of events and/or observation time cause the entire record to be ignored.

Finally the user is prompted for the name of the file to which output is to be written.

Optionally, the program will print out the first five records of the DATA file read, so that the user can verify that the correct values have been entered. Summary statistics (mean, min, max, number missing) are printed for each of the variables, and these can also serve as a useful check.

3.2 Step 2: Model specification

The user is prompted to specify the model to be used in the analysis. The number and names of the regressor variables to be included are entered. In addition, as long as two or more regressor variables are specified, the user is asked which 2-way interaction terms, if any, between the variables

are to be included in the model. For each interaction a new covariate is formed by multiplying together the two corresponding regressor variable values for each subject in turn. Any subject who has a missing value for one or more of the included covariates will be excluded from the analysis of this model. The value of p as defined in Section 2, is then the number of variables included in the model plus the number of interactions specified (if any).

To aid in the specification of the model, at any stage all previous models fitted are listed. The user can choose any of these (including the “null” model in which no covariates are included) as a “base” model and only the additional covariates and interactions that are in the desired model but not in the base model need be specified. All covariates in the base model are automatically included.

3.3 Likelihood maximization

The NAG [13] subroutine E04UCF, which uses a sequential quadratic programming method for optimizing a nonlinear function, is employed to maximize the log-likelihood \mathcal{L} . The subroutine may be replaced by a Newton-Raphson or other similar algorithm; however the authors have had good experience with the one recommended. (The older NAG routine E04VDF can also be substituted.) Initially, starting values for μ and ν are obtained by the method of moments, with $\beta = \mathbf{0}$. If, for any reason, the algorithm fails to converge to the optimum, an appropriate error message is displayed. In our experience with several data sets, large and small, we have not encountered any problems with non-convergence. (The program does however impose working lower bounds of 10^{-6} on the values of μ and ν which helps avoid possible problems.)

3.4 Step 4: Output of parameter estimates, standard errors and test statistics

After successful convergence, the program prints out the following results both to the screen and to the designated output file:

- The maximum likelihood estimates (MLEs), $\hat{\nu}$ and $\hat{\mu}$, of the shape parameter ν and the mean $\mu = \nu\gamma$ of the mixing distribution along with their standard errors.
- The deviance, i.e. $-2\mathcal{L}(\hat{\mu}, \hat{\nu}, \hat{\beta})$. For comparison, the deviance of the null model (i.e. that without any covariates $\beta = \mathbf{0}$) is also displayed.

- The likelihood ratio statistic which is the difference of the two deviances in the above item. The degrees of freedom are equal to p , the number of covariates fitted (including any interaction terms). By comparison with the percentage points of the chi-squared distribution with p degrees of freedom, the overall significance of the fitted covariates can be assessed. Similarly if two models have been fitted, with p_1 and p_2 degrees of freedom ($p_1 < p_2$) with one nested within the other, the significance of the improvement in fit by including the extra covariates in the second model can be assessed by comparing the difference in the respective deviances with percentage points of the chi-squared distribution with $p_2 - p_1$ degrees of freedom.
- The score test statistic. This can be used in a similar fashion as the likelihood ratio statistic to assess the overall significance of the fitted covariates.
- For each covariate (and interaction) included in the model, the program displays the estimated regression coefficient, its standard error, its standardized Z-value and a 95% confidence interval for it.
- Finally, the estimated correlation matrix for the parameter estimates $(\hat{\mu}, \hat{\nu}, \hat{\beta})$ is written to the output file only. This is obtained from the covariance matrix which is the inverse of the information matrix as explained in Section 2.

The significance tests and confidence intervals are based on the asymptotic normal approximation. The user is then asked if he wants fit a new model with added covariates or interactions. If yes, a list of the last 10 models previously fitted with corresponding deviance values are printed to aid in the choice of variables to include. Steps 2–4 are then repeated.

4 Sample runs

Sample runs for two data sets are described. The first is a rather simple one with a small sample size, a common followup time for all subjects and a single binary covariate. The second data set is a very large one with many covariates and varying followup times. The terminal sessions for the PC version are given in Appendices A and B, respectively. Sessions for the UNIX version are almost identical.

4.1 Mammary tumor rat data

Gail, Santner and Brown [4] presented data on times to mammary tumor for 23 female rats on retinyl acetate treatment and 25 female control rats. The DICTIONary file appears in Table 1 and the DATA file in Table 2. This data set has become a standard one and has been used by several authors, including most recently Sinha [14]. Randomization occurred on day 60 of the study and the study ended on day 182. Thus the followup time T_i for each rat was 122 days. Gail et al. [4] give the actual times of each incident mammary tumor, but we only use the number of tumors occurring between days 60 and 182. There is only one covariate ($p = 1$), namely treatment ($x = 1$ for retinyl, $x = 0$ for control). The sample run is displayed in Appendix A. The null model omitting the treatment covariate is fitted first, then the full model is fitted, using the covariate. It can be seen from the output that the effect of the retinoid treatment is to significantly reduce the incidence of mammary tumors.

4.2 Skin cancer prevention clinical trial data

Clark et al. [15] describe an ongoing double blind multi-center randomized clinical trial to test the benefits of a nutritional supplement of selenium for the prevention of recurrent skin cancer. The trial has also been described by Abu-Libdeh et al. [3]. Starting in 1983, entering patients have been randomized to either selenium (Se) supplementation or to placebo. We use interim data from this trial to illustrate our program. There are 1277 patients. The response from each patient consists of a list of times (measured from date of entry) at which new squamous cell carcinomas (SCCs) were diagnosed. The length of followup varies from patient to patient – the average followup being about five years. Besides treatment assignment, a number of baseline covariates were collected from each patient at the time of entry. These included plasma selenium level in mg/dl (Se1), age, gender, weight (in kg.), sun damage, previous number of SCCs diagnosed **before** entry (#Prev_SCC), and the attending clinic. A more complete description is given in Abu-Libdeh et al. (1990, Sec.4) – see also the “Quick Data Summary” table which is part of the output from the sample run in Appendix B. The null model, without covariates, is first run. We next ran a model with covariate “baseline Se level”, next a model with “age” and “gender” added; finally a model which additionally included an interaction between age and gender. We see that baseline plasma selenium level and gender are significant. Increased selenium level is protective, and males are at higher risk. (Note here that a

principal use of the program would be to assess the significance of the covariate “treatment” after adjusting for the important other covariates. However, because this is only an interim analysis of an ongoing blinded trial, this cannot be done here – see the article by Green, Fleming and O’Fallon [16] on interim reporting of results from clinical trials.)

5 Hardware and Software Specifications

The program was written using FORTRAN-77 for a SUN SPARCstation [17]. CPU time for the example of Section 4.1 was 0.23 secs on a SPARC 10, 0.63 secs on a SPARC 2, and 1.47 secs on a SPARC 1. The corresponding CPU times for the example of Section 4.2, with its multiple models fitted, were 22.07, 77.37 and 190.77 seconds. A Microsoft FORTRAN PowerStation [18] program has also been compiled to run on a 386 or 486 IBM PC or compatible.

6 Mode of Availability of the Program

Copies of the program are available upon request from the corresponding author, Bruce W. Turnbull.

7 Acknowledgement

This research was supported principally by grant R01 CA61120, also by grants R01 GM28364 and R01 CA 49764, all from the U.S. National Institutes of Health.

Appendix A. Sample terminal session – Mammary tumor rat data

MIXED EFFECTS POISSON REGRESSION MODEL FOR REPEATED EVENTS

WHAT IS THE DICTIONARY FILE? rat.dct

SPECIFY FILE FOR PRINTED OUTPUT: rat.out

READING INPUT DATA...PLEASE WAIT

48 RECORDS ARE BEING READ

0 RECORDS ARE BEING IGNORED DUE TO MISSING NUMBER OF EVENTS

AND/OR EXPOSURE TIME

PRINT FIRST FIVE RECORDS? (Y/N): y

RECORD 1

 1.0000 122.0000 1.0000

RECORD 2

 0.0000 122.0000 1.0000

RECORD 3

 2.0000 122.0000 1.0000

RECORD 4

 1.0000 122.0000 1.0000

RECORD 5

 4.0000 122.0000 1.0000

QUICK DATA SUMMARY:	MEAN	MIN	MAX	# MISSING
N_Tumors	4.417	.0	13.0	0.
Exp_Time	122.000	122.0	122.0	0.
TREATMENT	1.521	1.0	2.0	0.

DO YOU WISH TO CONTINUE? (Y/N) y

MAXIMIZING THE LIKELIHOOD FUNCTION FOR THE MODEL WITHOUT ANY COVARIATES
(NULL MODEL)....PLEASE WAIT

SUCCESSFUL CONVERGENCE HAS BEEN REACHED AFTER 15 ITERATIONS

STATISTICAL OUTPUT BASED ON 48 RECORDS

PARAMETERS OF THE MIXING DISTRIBUTION

SHAPE = 2.29555 MEAN = 0.036202
ST.DEV = (0.73475) ST.DEV = (0.004252)

DEVIANC FOR NULL MODEL (-2*(LOG-LIK)) = 1794.62442

DO YOU WISH TO ADD COVARIATES TO THE MODEL? (Y/N) y

INDEXED LIST OF COVARIATES

1: TREATMENT

HOW MANY COVARIATES DO YOU WISH TO USE? (1 -- 1) 1

ENTER INDEX OF CHOSEN COVARIATES. 1

-----OPTIMIZING THE LIKELIHOOD FUNCTION-----

SUCCESSFUL CONVERGENCE HAS BEEN REACHED AFTER 15 ITERATIONS

STATISTICAL OUTPUT BASED ON 48 RECORDS

PARAMETERS OF THE MIXING DISTRIBUTION

SHAPE = 3.75301 MEAN = 0.009546
ST.DEV = (1.51572) ST.DEV = (0.003431)

DEVIANC: CURRENT MODEL = 1781.46912 NULL = 1794.62442

LIK. RATIO TEST STATISTIC = 13.15530 WITH 1 DEGREE(S) OF FREEDOM.

SCORE TEST STATISTIC = 24.02362 WITH 1 DEGREE(S) OF FREEDOM.

COVARIATE	M.L.E.	ST.DEV.	Z-VALUE	95% C. INTERVAL
TREATMENT	0.823024*	0.212742	3.8687	(0.406 , 1.240)

*: SIGNIFICANT AT THE 5% LEVEL

DO YOU WISH TO ADD COVARIATES TO THE MODEL? (Y/N) n

Appendix B. Sample terminal session – Clinical trial data

MIXED EFFECTS POISSON REGRESSION MODEL FOR REPEATED EVENTS

WHAT IS THE DICTIONARY FILE? skin.dct

SPECIFY FILE FOR PRINTED OUTPUT: skin.out

READING INPUT DATA...PLEASE WAIT

1277 RECORDS ARE BEING READ.

0 RECORDS ARE BEING IGNORED DUE TO MISSING NUMBER OF EVENTS

AND/OR EXPOSURE TIME

PRINT FIRST FIVE RECORDS? (Y/N): y

RECORD 1

0.0000	2166.0000	2.0000	0.0000	1.0000	0.0000	65.9000
1.0000	4.0000	6.0000	0.0000	5.0000	122.0000	3.0000
0.0000	0.0000	1.0000				

RECORD 2

0.0000	169.0000	2.0000	1.0000	2.0000	0.0000	62.3000
1.0000	4.0000	5.0000	1.0000	7.5000	104.0000	2.0000
0.0000	1.0000	1.0000				

RECORD 3

0.0000	967.0000	1.0000	1.0000	1.0000	1.0000	90.9000
0.0000	6.0000	5.0000	0.0000	6.0000	84.0000	0.0000
1.0000	1.0000	1.0000				

RECORD 4

0.0000	2061.0000	2.0000	1.0000	1.0000	0.0000	102.3000
0.0000	5.0000	5.0000	0.0000	5.5000	107.0000	0.0000
1.0000	1.0000	1.0000				

RECORD 5

0.0000	1716.0000	2.0000	0.0000	1.0000	0.0000	70.9000
0.0000	4.0000	5.0000	0.0000	5.5000	101.0000	11.0000
0.0000	0.0000	1.0000				

QUICK DATA SUMMARY: MEAN MIN MAX # MISSING

N_Tumors	.538	.0	13.0	0.
Exp_Time	1593.545	1.0	3276.0	0.
T_Group	1.501	1.0	2.0	0.
AGE>64	.525	.0	1.0	0.
GENDER M=1	1.255	1.0	2.0	0.
Cur_Smoker	0.284	.0	1.0	0.
Weight_Kg	77.338	40.5	143.2	1.
Farm_GE_15	0.344	.0	1.0	0.
Sunburn(a)	4.741	4.0	6.0	0.
Suntan(b)	4.836	4.0	6.0	0.
Sunscrn<5	0.282	.0	1.0	0.
Sun_Damage	4.733	1.0	9.0	0.
Se1	114.243	45.0	220.0	0.
#Prev_BCE	2.898	.0	66.0	0.
#Prev_SCC	0.979	.0	16.0	0.
In_Patient	0.458	.0	1.0	0.
Clinic_#	3.165	1.0	8.0	0.

DO YOU WISH TO CONTINUE? (Y/N) y

MAXIMIZING THE LIKELIHOOD FUNCTION FOR THE MODEL WITHOUT ANY COVARIATES

(NULL MODEL).....PLEASE WAIT

SUCCESSFUL CONVERGENCE HAS BEEN REACHED AFTER 17 ITERATIONS

STATISTICAL OUTPUT BASED ON 1277 RECORDS

PARAMETERS OF THE MIXING DISTRIBUTION

SHAPE = 0.36653	MEAN = 0.000342
ST.DEV = (0.03779)	ST.DEV = (0.000021)

DEVIANCE FOR NULL MODEL (-2*(LOG-LIK)) = 11880.57057

DO YOU WISH TO ADD COVARIATES TO THE MODEL? (Y/N) y

INDEXED LIST OF COVARIATES

1: T_Group	2: AGE>64	3: GENDER M=1	4: Cur_Smoker	5: Weight_Kg
6: Farm_GE_15	7: Sunburn(a)	8: Suntan(b)	9: Sunscrn<5	10: Sun_Damage
11: Se1	12: #Prev_BCE	13: #Prev_SCC	14: In_Patient	15: Clinic_#

HOW MANY COVARIATES DO YOU WISH TO USE? (1 -- 15) 1

ENTER INDEX OF CHOSEN COVARIATES: 11

-----OPTIMIZING THE LIKELIHOOD FUNCTION----- (PLEASE WAIT)

SUCCESSFUL CONVERGENCE HAS BEEN REACHED AFTER 30 ITERATIONS

STATISTICAL OUTPUT BASED ON 1277 RECORDS

PARAMETERS OF THE MIXING DISTRIBUTION

SHAPE = 0.37118	MEAN = 0.000646
ST.DEV = (0.03844)	ST.DEV = (0.000209)
DEVIANCE: CURRENT MODEL =11876.57471	NULL =11880.57057
LIK. RATIO TEST STATISTIC = 3.99586	WITH 1 DEGREE(S) OF FREEDOM.
SCORE TEST STATISTIC = 4.03472	WITH 1 DEGREE(S) OF FREEDOM.

COVARIATE	M.L.E.	ST.DEV.	Z-VALUE	95% C. INTERVAL
Se1	-0.005638*	0.002823	-1.9967	(-0.011 , 0.000)

*: SIGNIFICANT AT THE 5% LEVEL

DO YOU WISH TO ADD COVARIATES TO THE MODEL? (Y/N) y

INDEXED LIST OF COVARIATES

1: T_Group	2: AGE>64	3: GENDER M=1	4: Cur_Smoker	5: Weight_Kg
6: Farm_GE_15	7: Sunburn(a)	8: Suntan(b)	9: Sunscrn<5	10: Sun_Damage
11: Se1	12: #Prev_BCE	13: #Prev_SCC	14: In_Patient	15: Clinic_#

PREVIOUS MODEL(S) FITTED

MDL 0: NULL MODEL

DEVIANCE: 11880.57057

MDL 1: Se1

DEVIANCE: 11876.57471

WHICH MODEL DO YOU WISH TO USE AS A BASE? 1
 HOW MANY NEW VARIABLES DO YOU WISH TO ADD? 2
 ENTER INDEX OF CHOSEN COVARIATES: 2 3
 HOW MANY NEW INTERACTIONS? (0=none) 0
 -----OPTIMIZING THE LIKELIHOOD FUNCTION----- (PLEASE WAIT)
 SUCCESSFUL CONVERGENCE HAS BEEN REACHED AFTER 44 ITERATIONS
 STATISTICAL OUTPUT BASED ON 1277 RECORDS
 PARAMETERS OF THE MIXING DISTRIBUTION
 SHAPE = 0.42968 MEAN = 0.001761
 ST.DEV = (0.04649) ST.DEV = (0.000715)
 DEVIANCE: CURRENT MODEL =11829.50071 NULL =11880.57057
 LIK. RATIO TEST STATISTIC = 51.06986 WITH 3 DEGREE(S) OF FREEDOM.
 SCORE TEST STATISTIC = 60.78872 WITH 3 DEGREE(S) OF FREEDOM.

COVARIATE	M.L.E.	ST.DEV.	Z-VALUE	95% C. INTERVAL
Se1	-0.007171*	0.002801	-2.5602	(-0.013 , -0.002)
AGE>64	0.440036*	0.122981	3.5781	(0.199 , 0.681)
GENDER M=1	-0.918275*	0.159514	-5.7567	(-1.231 , -0.606)

*: SIGNIFICANT AT THE 5% LEVEL

DO YOU WISH TO ADD COVARIATES TO THE MODEL? (Y/N) y

INDEXED LIST OF COVARIATES

1: T_Group	2: AGE>64	3: GENDER M=1	4: Cur_Smoker	5: Weight_Kg
6: Farm_GE_15	7: Sunburn(a)	8: Suntan(b)	9: Sunscrn<5	10: Sun_Damage
11: Se1	12: #Prev_BCE	13: #Prev_SCC	14: In_Patient	15: Clinic_#

PREVIOUS MODEL(S) FITTED

MDL 0: NULL MODEL

DEVIANCE: 11880.57057

MDL 1: Se1

DEVIANCE: 11876.57471

MDL 2: Se1 AGE>64 GENDER M=1

DEVIANCE: 11829.50071

WHICH MODEL DO YOU WISH TO USE AS A BASE? 2

HOW MANY NEW VARIABLES DO YOU WISH TO ADD? 0

HOW MANY NEW INTERACTIONS? (0=none) 1

ENTER PAIR OF INDICES FOR ONE INTERACTION 2 3

-----OPTIMIZING THE LIKELIHOOD FUNCTION----- (PLEASE WAIT)

SUCCESSFUL CONVERGENCE HAS BEEN REACHED AFTER 45 ITERATIONS

STATISTICAL OUTPUT BASED ON 1277 RECORDS

PARAMETERS OF THE MIXING DISTRIBUTION

SHAPE = 0.43065 MEAN = 0.002174

ST.DEV = (0.04662) ST.DEV = (0.000999)

DEVIANCE: CURRENT MODEL =11828.51503 NULL =11880.57057

LIK. RATIO TEST STATISTIC = 52.05554 WITH 4 DEGREE(S) OF FREEDOM.

SCORE TEST STATISTIC = 62.55758 WITH 4 DEGREE(S) OF FREEDOM.

COVARIATE	M.L.E.	ST.DEV.	Z-VALUE	95% C. INTERVAL
Se1	-0.007165*	0.002801	-2.5582	(-0.013 , -0.002)
AGE>64	0.065689	0.397167	0.1654	(-0.713 , 0.844)
GENDER M=1	-1.096237*	0.242267	-4.5249	(-1.571 , -0.621)
AGE>*GEND	0.316835	0.320077	0.9899	(-0.311 , 0.944)

*: SIGNIFICANT AT THE 5% LEVEL

DO YOU WISH TO ADD COVARIATES TO THE MODEL? (Y/N) n

References

- [1] H. Abu-Libdeh, B.W. Turnbull and L.C. Clark, Analysis of multi-type recurrent events in longitudinal studies; application to a skin cancer prevention trial, *Biometrics* 46 (1990) 1017-1034.
- [2] P.F. Thall, Mixed Poisson likelihood regression models for longitudinal interval count data, *Biometrics* 44 (1988) 197-209.
- [3] L. Freedman, R. Sylvester and D.P. Byar, Using permutation tests and bootstrap confidence limits to analyze repeated events data from clinical trials, *Controlled Clinical Trials* 10, 129-141.
- [4] M.H. Gail, T.J. Santner and C.C. Brown, An analysis of comparative carcinogenesis experiments based on multiple times to tumor, *Biometrics* 36 (1980) 255-266.
- [5] H. Ascher and H. Feingold, *Repairable Systems Reliability* (Marcel Dekker, New York, NY, 1984).
- [6] P.D. Allison, *Event History Analysis: Regression for Longitudinal Data* (Sage Publications, Beverly Hills, CA, 1984).
- [7] J.F. Lawless, *Statistical Models and Methods for Lifetime Data* (Wiley, New York, NY, 1982).
- [8] J.F. Lawless, Regression methods for Poisson process data, *Journal of the American Statistical Association* 82 (1987) 808-815.
- [9] N.L. Johnson and S. Kotz, *Distributions in Statistics: Continuous Univariate Distributions I* (Houghton Mifflin, Boston, MA, 1970).
- [10] J.F. Lawless, Negative binomial regression models, *Canadian Journal of Statistics* 15 (1987) 209-225.
- [11] C.R. Rao, *Linear Statistical Inference and its Applications* (Wiley, New York, NY, 1973)
- [12] D.R. Cox and D.V. Hinkley, *Theoretical Statistics* (Chapman and Hall, London, UK, 1974).

- [13] The Numerical Algorithms Group, The NAG Foundation Library Handbook (NAG, Oxford UK, 1992)
- [14] Sinha, D., Semiparametric Bayesian analysis of multiple event time data, Journal of the American Statistical Association 88 (1993) 979-983.
- [15] L.C. Clark, B.H. Patterson, D.L. Weed, and B.W. Turnbull, Design issues in cancer chemo-prevention trials using micronutrients: application to skin cancer, Cancer Bulletin 43 (1991) 519-524.
- [16] S.J. Green, T.R. Fleming and J.R. O'Fallon, Policies for study monitoring and interim reporting of results, Journal of Clinical Oncology 5 (1987) 1477-1484.
- [17] Sun Microsystems, Inc., Mountain View, California.
- [18] Microsoft FORTRAN PowerStation, Microsoft Corp., Redmond, Washington.

CAPTIONS FOR TABLES AND FIGURE

Table 1: Dictionary File (rat.dct) for Mammary Tumor Rat Example

Table 2: Data File (rat.dat) for Mammary Tumor Rat Example

Figure 1: Flow Chart

Table 1: Dictionary File (rat.dct) for Mammary Tumor Rat Example

```
rat.dat
2 48
ID
N_Tumors
Exp_Time
TREATMENT
```

Table 2: Data File (rat.dat) for Mammary Tumor Rat Example

ID	No. of tumors	Observation time	Treatment group
1	1	122	1
2	0	122	1
3	2	122	1
4	1	122	1
5	4	122	1
6	3	122	1
7	6	122	1
8	1	122	1
9	1	122	1
10	5	122	1
11	2	122	1
12	1	122	1
13	5	122	1
14	2	122	1
15	3	122	1
16	4	122	1
17	5	122	1
18	5	122	1
19	1	122	1
20	2	122	1
21	6	122	1
22	0	122	1
23	1	122	1
24	7	122	2
25	11	122	2
26	9	122	2
27	2	122	2
28	9	122	2
29	4	122	2
30	6	122	2
31	7	122	2
32	6	122	2
33	1	122	2
34	13	122	2
35	2	122	2
36	1	122	2
37	10	122	2
38	4	122	2
39	5	122	2
40	11	122	2
41	11	122	2
42	9	122	2
43	12	122	2
44	1	122	2
45	3	122	2
46	1	122	2
47	3	122	2
48	3	122	2

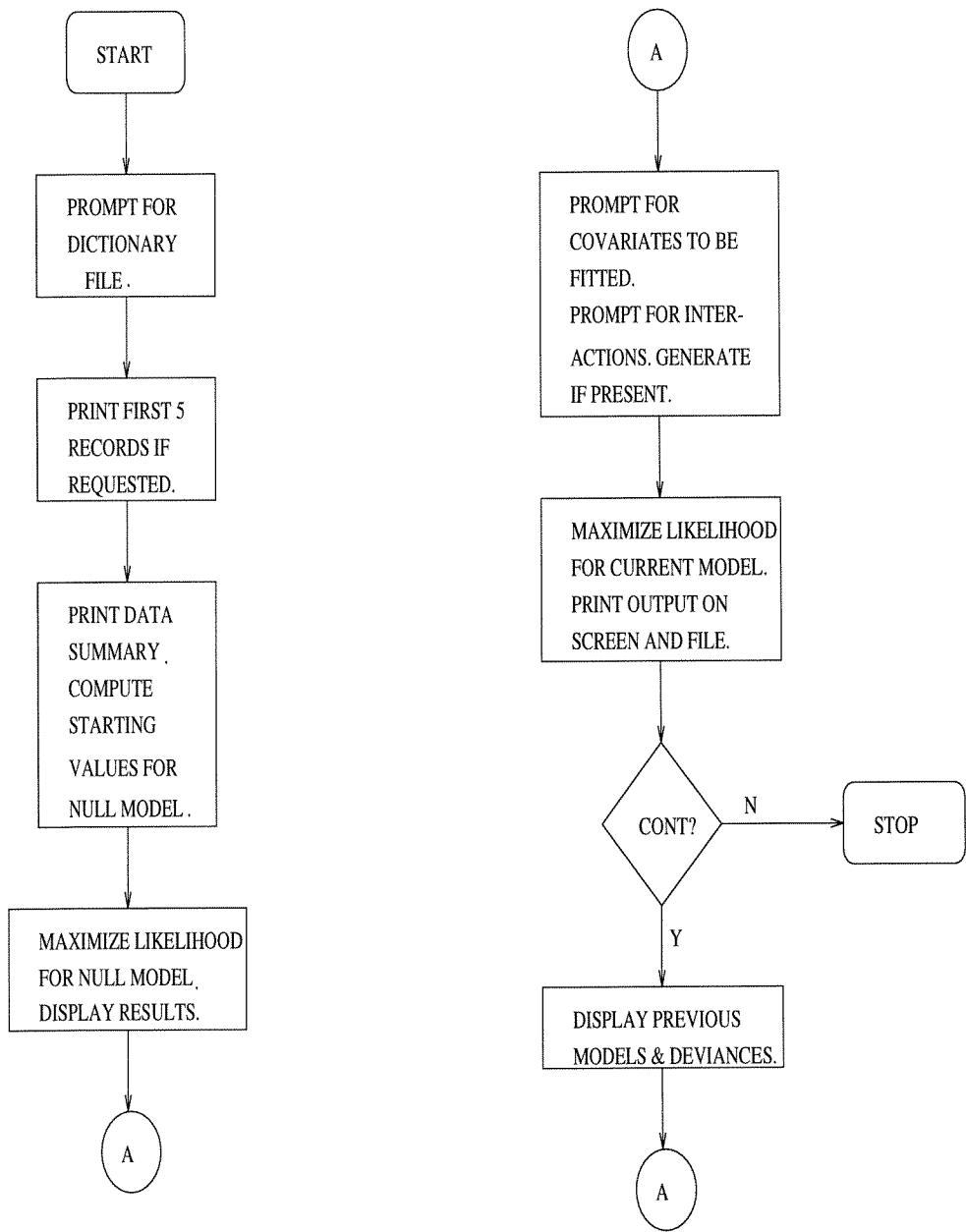


Figure 1: Flow Chart for the Program

program

Thu Jun 9 13:42:25 1994

1

```
C * ***** MIXED POISSON REGRESSION MODEL FOR *****
C * ***** REPEATED EVENT DATA *****
C
```

```
C THE PROGRAM CALCULATES:
C 1. MLE'S AND SDS OF MIXING DISTRIBUTION PAMS
C 2. MLE'S AND SDS OF REGRESSION PAMS
C 3. Z-VALUES FOR REGRESSION PAMS
```

```
C 4. A NORMAL-BASED 95% CI FOR EACH REGRESSION PARM
C 5. L.R. AND SCORE TEST STATISTICS FOR REGRESSION PAMS.
C 6. CORRELATION MATRIX.
```

```
C SUBROUTINES CALLED IN THE PROGRAM:
```

```
C PARTIII: EVALUATES SCORE VECTOR AND INFO-MATRIX
C          SCORE: EVALUATES SCORE STATISTIC FOR H(0): REG. COEFFS = ZERO.
C          ARINV: INVERTS HESSIAN MATRIX.
C          E04VDF: THIS NAG ROUTINE MAXIMIZES THE LOG-LIKELIHOOD FUNCTION.
C          E04VDM: THIS NAG SUBROUTINE IS CALLED BY E04VDF.
C          OBJSFUN: THIS SUBROUTINE COMPUTES THE OBJECTIVE FUNCTION FOR THE
C          NAG OPTIMIZATION FUNCTION FOR THE CURRENT MODEL.
C          OBNULL: THIS SUBROUTINE COMPUTES THE OBJECTIVE FUNCTION FOR THE
C          NAG OPTIMIZATION FUNCTION FOR THE NULL MODEL.
```

```
C DEFINITIONS OF IMPORTANT PROGRAM VARIABLES
```

```
C A(, ) : NEGATIVE OF SAMPLE INFORMATION MATRIX.
C ALERT : 1 IF NUMBER OF SUBJECTS HAS BEEN REDUCED DUE TO
C          SOME MISSING OBSERVATION
C ANS : CHARACTER VARIABLE THAT INDICATES WHETHER USER
          USER WANTS TO CONTINUE
```

```
C BASE : BASE MODEL THAT USER WANTS TO BUILD ON
C BASENULL( ) : CONTAINS MLE OF BASE MODEL.
C C1,C2 : CONTAIN INFORMATION ON NUMBER OF VARIABLES IN BASE MODEL
C CHI : CHI SQUARE STATISTIC FOR TESTING SIGNIFICANCE OF
```

```
C COL(25) : OF THE REGRESSION
C COVNAME() : NAMES OF COVARIATES IN DATA FILE
C COVRT( ) : NAMES OF COVARIATES AVAILABLE FOR ANALYSIS.
C COVTEMP : NUMBER OF COVARIATES ENTERING ANALYSIS
```

```
C DICTION : DICTIONARY FILE NAME
C DLGLK : DEVIANCE FOR NULL MODEL
C EPS : RADIUS OF CONVERGENCE
C FLAG( ) : 1 IF ANY MISSING ENTRIES FOR ANY COVARIATE
```

```
C GMA(3) : NAMES OF BASELINE PARAMETERS.
C GMEAN : MEAN NUMBER OF TUMORS (USED IN MOM ESTIMATION
C ICOL : > 0 IF ANY MODEL OTHER THAN THE NULL HAS BEEN FITTED
C ID : PATIENT DESCRIPTOR
```

```
C INPUTF : FILE CONTAINING COVARIATE DATA.
C INTERD( ) : INDICES OF VARIABLES APPEARING IN INTERACTIONS
C LIKRATIO : ARRAY OF DEVIANCES OF PREVIOUS MODELS
C M : NUMBER OF PARAMETERS TO BE ESTIMATED.
C MAT( , ) : WORKING MATRIX FOR THE SUBROUTINE ARINV.
C MAXX( ) : ARRAY OF MAXIMUM VALUES OF EACH COVARIATE
```

```
C MDL : MODEL CHOICE POINTER (=2)
C MINN( ) : ARRAY OF MINIMUM OF EACH COVARIATE
C MISSING( ) : ARRAY OF NUMBER MISSING FOR EACH COVARIATE
C MODINT( , ) : ARRAY OF PREVIOUS INTERACTIONS FITTED
C MODVAR( , ) : ARRAY OF PREVIOUS COVARIATES FITTED
C : THE SIZE OF MOVAR AND MODINT LIMITS ONE TO 25 MODELS.
```

```
C NCOV : NUMBER OF COVARIATES ENTERING ANALYSIS.
C NINTS : NUMBER OF PAIRWISE-INTERACTIONS IN THE ANALYSIS
C NRUN : NUMBER OF MODELS ALLOWED IN ONE SESSION (25)
```

```
C NUMMISS : NUMBER MISSING IN ANY ANALYSIS
C NVARS : NUMBER OF COVARIATES IN THE DATA FILE INPUT
C OBJF : VALUE OF LOG-LIKELIHOOD FUNCTION.
C P( , ) : WORKING MATRIX
C PARMBASE( ) : STORAGE OF PARM ESTIMATES OF MIXING DIST
C : COEFFICIENTS FOR PAST MODELS
```

```
C PARMCOV( , ) : STORAGE OF PARM ESTIMATES OF REGRESSION
```

```
C : COEFFICIENTS FOR PAST MODELS
C NSUBS : NUMBER OF OBSERVATIONS ENTERING ANALYSIS.
C OUTPUT : OUTPUT FILE NAME FOR ANALYSIS RESULTS.
C ST( ) : SQUARE ROOTS OF DIAGONAL ELEMENTS OF COV. MATRIX
C STAT : TEMPORARY STORAGE FOR BASELINE STARTING VALUES.
C START(5) : TEMPORARY STORAGE FOR BASELINE STARTING VALUES.
C SUM() : SUM OF EACH COVARIATE
C TOTAL : TOTAL NUMBER OF TUMORS (USED IN MOM ESTIMATION)
C X0() : VECTOR OF PARAMETER ESTIMATES AT ANY ITERATION.
C XINC( ) : NEGATIVE OF THE SCORE VECTOR.
C XX( , ) : INPUT COVARIATE MATRIX OF ALL COVARIATES
C XX1( , ) : COVARIATES ENTERING ANALYSIS.
```

```
DOUBLE PRECISION ST(25),ID,STAT,MAT(25,25),CHI
DOUBLE PRECISION PARMCOV(25,25),PARMBASE(25,3),BASENULL(3)
DOUBLE PRECISION A(25,25),XO(25),XINC(25),X(25)
DOUBLE PRECISION MAXX(25),MINN(25),SUM(25),MISSING(25)
DOUBLE PRECISION LIKRATIO(25),DLGLK
DOUBLE PRECISION TOTAL,GMEAN,VAR
DOUBLE PRECISION RESP(3000,2)
INTEGER M,I,J,ICOL,NVARS
INTEGER FLAG(1ALLOCATABLE)(:)
INTEGER COL(25),BASE,NINTS,C1,C2,NUMMISS
NRUN,MODVAR(25,25),MODINT(25,25)
INTEGER INTERDEX(10,10)
COMMON XX(3000,30),XX1(3000,30),P(25,25),
      + MDL,NSUBS,START(5),COVTEMP
```

```
CHARACTER*12 COVNAME(100)
CHARACTER*10 N_EVENT, EXP_TIME
CHARACTER*10 COVRT(25)*10,GMA(3)*5,ANS
LOGICAL EX
```

```
C THE FOLLOWING DEFINITIONS ARE FOR USING THE OPTIMIZATION
C ROUTINES IN NAG. SEE NAG MANUAL FOR A DESCRIPTION OF THESE
C VARIABLES.
```

```
INTEGER N,NCLIN,NCNLN
INTEGER NRWQ,NROWQ,NROWR
INTEGER LIWORK,LWORK
DOUBLE PRECISION BIGND,OBFE
INTEGER IFAIL,ITER
DOUBLE PRECISION AA(1,25),BI(25),BU(25),C(1),CJAC(1,25),
      + CLAMDA(25),R(25,25),
      + WORK(2000)
+
```

```

DOUBLE PRECISION USER(1)
DOUBLE PRECISION OBJGRD(25)
INTEGER ISTATE(25),IWORK(100),IUSER(1)
EXTERNAL OBFJFUN,E04UFC,E04UDM,E04UEF,OBNNULL
DATA NRUN,NINTS,M,NCOV/0,0,0,0/

C SET ALL THE VARIABLES THAT WILL BE USED IN THE NAG SUBROUTINE
NRWR=25
NCLIN=0
NCNLIN=0
NROWA=1
NROWJ=1
L1WORK=1.00
LWORK=2000
BIGBND=.0E10

C SET LOWER BOUNDARIES FOR SHAPE AND MEAN OF GAMMA DIST TO 0
BL(1)=0.000001
BL(2)=0.000001

C INITIALIZE ARRAYS AND VARIABLES THAT WILL BE USED IN THE PROGRAM
DO 1994 I = 1,3000
    RESP(I,1)=0
    RESP(I,2)=0
    DO 1994 J = 1,30
        XX(I,J) = 0
        XX1(I,J) = 0
    1994
    DO 1996 I = 1,25
        MODVAR(I,J) = 0
        MODINT(I,J) = 0
        PARMCOV(I,J) = 0
        P(I,J) = 0
    1996

C PRINT HEADER
      WRITE(*,1024)
      WRITE(*,72)
      72 FORMAT(5X,'MIXED EFFECTS POISSON REGRESSION MODEL FOR',1X,
      + 'REPATED EVENTS')
      WRITE(*,1024)
      WRITE(*,*)
      WRITE(*,*)

C PROMPT FOR DICTIONARY AND OUTPUT FILE
      222 WRITE(*,8010)
      8010 FORMAT(1X,'WHAT IS THE DICTIONARY FILE? ',$:)
      READ(*,'(A)') DICTION
      INQUIRE(FILE=DICTION,EXIST=EX)
      IF (EX) THEN
          OPEN(26,FILE=DICTION,STATUS='OLD')
      ELSE
          WRITE(*,*) 'NO SUCH FILE!'
          GOTO 222
      ENDIF
      WRITE(*,*)

      WRITE(*,8011)
      8011 FORMAT(1X,'SPECIFY FILE FOR PRINTED OUTPUT: ',$:)
      READ(*,'(A)') OUTPUTF
      OPEN(25,FILE=OUTPUTF)

      READ(26,'(A)') INPUTF
      INQUIRE(FILE=INPUTF,EXIST=EX)
      IF (EX) THEN
          OPEN(28,FILE=INPUTF,STATUS='OLD')
      ENDIF

C READ RELEVANT INFORMATION FROM DICTIONARY FILE AND DATA FILE
      READ(26,*)
      READ(26,'(A)') COVNAME(1),N_EVENT, EXP_TIME,
      + (COVNAME(J), J = 2, NVARS)
      CLOSE(26)

C THE DATA FILE HAS ID IN THE FIRST COLUMN
      NVARS = NVARS-1
      ALLOCATE(FLAG(NSUBS))

      WRITE(*,*)
      WRITE(*,*) 'READING INPUT DATA .. PLEASE WAIT'
      WRITE(*,*)

```

Program Thu Jun 9 13:42:25 1994

3
Thu Jun 9 13:42:25 1994

```

Thu Jun 9 13:42:25 1994

        WRITE(*,*), 'EXITING THE PROGRAM'
        GOTO 9
ENDIF

GMEAN=TOTAL/NSUBS

DO 7071 I = 1, NSUBS
    VAR = VAR + (XX(I,1) - GMEAN) * (XX(I,1) - GMEAN)
    IF (IFAIL .EQ. 4) THEN
        WRITE(*,*), 'ITERATION LIMIT EXCEEDED!'
        WRITE(*,*), 'EXITING THE PROGRAM'
        GOTO 9
    ENDIF

    XO(1) = (VAR/TOTAL) - 1.0
    XO(2) = GMEAN
    M=2
    N=M

    C SET ERROR FLAG FOR NAG EQUAL TO -1. SEE MANUAL FOR AN EXPLANATION.
    IFAIL=-1
    COVTEMP=0

    WRITE(*,*)
    WRITE(*,7075)
    WRITE(*,*), '(NULL MODEL).....PLEASE WAIT'
    7075 FORMAT(1X,'MAXIMIZING THE LIKELIHOOD FUNCTION FOR THE',1X,
    + 'MODEL WITHOUT ANY COVARIATES')
    CALL E04UEF('MAJOR PRINT LEVEL = 0')
    CALL E04UDM,OBJNULL,IITER,ISTATE,C,Cjac,CLAMDA,OBJF,OBJGRD,R,
    + EO4UDM,OBJNULL,IITER,ISTATE,C,Cjac,CLAMDA,OBJF,OBJGRD,R,
    + XO,IWORK,LWORK,WORK,LWORK,IUSER,USER,IFAIL)

    C PRINT ERROR MESSAGES DEPENDING ON THE VALUE OF IFAIL
    IF (IFAIL .EQ. 0) THEN
        WRITE(*,*)
        WRITE(*,6666) ITER
        WRITE(*,*)
    ENDIF

    6666 FORMAT(1X,'SUCCESSFUL CONVERGENCE HAS BEEN REACHED AFTER ',I4,
    + ' 1X, 'ITERATIONS')

    IF (IFAIL .EQ. 1) THEN
        WRITE(*,*)
        WRITE(*,6971)
        WRITE(*,*), 'EXITING THE PROGRAM'
        GOTO 9
    ENDIF

    IF (IFAIL .EQ. 2) THEN
        WRITE(*,*)
        WRITE(*,6972)
        WRITE(*,*), 'EXITING THE PROGRAM'
        GOTO 9
    ENDIF

    IF (IFAIL .EQ. 3) THEN
        WRITE(*,6973)

```

WRITE(*,*), 'EXITING THE PROGRAM'

```

        GOTO 9
    ENDIF

    IF (IFAIL .EQ. 6) THEN
        WRITE(*,6975)
        WRITE(*,*), 'EXITING THE PROGRAM'
        GOTO 9
    ENDIF

    IF (IFAIL .EQ. 7) THEN
        WRITE(*,6976)
        WRITE(*,*), 'EXITING THE PROGRAM'
        GOTO 9
    ENDIF

    IF (IFAIL .EQ. 9) THEN
        WRITE(*,6977)
        WRITE(*,*), 'EXITING THE PROGRAM'
        GOTO 9
    ENDIF

    C COMPUTE STANDARD ERRORS FOR MIXING DISTRIBUTION PARAMETERS
    CALL PARTIII(XO,A,M,XINC,MAT,NCOV,OBJF)
    CALL ARINV(M,MAT)
    NCOV=0

    DO 7095 I = 1,M
        IF (MAT(I,I) .LE. 0.0) THEN
            WRITE(*,*), 'INFORMATION MATRIX IS SINGULAR'
            GO TO 7078
        ENDIF
        ST(I)=SORT(MAT(I,I))
    7095

    C WRITE TO THE OUTPUT FILE THE MLE AND SDS OF THE PARMS OF THE MIXING DISTN
    WRITE(25,8005) NSUBS
    WRITE(25,*)
    WRITE(25,*)
    WRITE(*,8005) NSUBS
    WRITE(*,*), 'STATISTICAL OUTPUT BASED ON', I6, ' RECORDS'
    8005 FORMAT(15X,'PARAMETERS OF THE MIXING DISTRIBUTION')
    8006 FORMAT(15X,'PARMS OF THE MIXING DISTRIBUTION')

```

```

      WRITE(*,*)' MDL 0: NULL MODEL'
      WRITE(*,7092) DLGLK
      WRITE(*,*)'
      WRITE(*,8006)
      WRITE(*,*)'
      WRITE(*,*)'
      WRITE(25,1015)XO(1),XO(2),ST(1),ST(2)
      WRITE(*,1015)XO(1),XO(2),ST(1),ST(2)
      WRITE(*,*)'
      WRITE(*,*)'
      DLGLK = -2*OBJF
      IF(NCOV.EQ.0) THEN
        WRITE(25,1017) DLGLK
        WRITE(*,1017) DLGLK
      ENDIF

C   STORE MIXING DIST MLE'S IN BASENULL()
      DO 7085 I = 1,M
        BASENULL(I)=XO(I)

      DO 7082 I = 1, NSUBS
        DO 7082 J = 1, NVARS+2
          XX1(I,J)=0
        WRITE(*, *)
      C   PROMPT FOR REGRESSION
        WRITE(*,7077)
        FORMAT(' DO YOU WISH TO ADD COVARIATES TO THE MODEL? (Y/N) ',,$)
        READ(*,'(A)') ANS
        IF((ANS.EQ.'N') .OR. (ANS.EQ.'n') .OR.
        + (ANS.EQ.'NO') .OR. (ANS.EQ.'no')) THEN
          GOTO 9
        ENDIF

        IF ((ANS.EQ.'Y') .OR. (ANS.EQ.'y')
        + .OR. (ANS.EQ.'Yes') .OR. (ANS.EQ.'YES')) THEN
          GOTO 7081
        ELSE
          WRITE(*,*) ' INVALID ANSWER '
          GOTO 7078
        ENDIF

C   SET UP COVARIATE SELECTION
      7081 WRITE(*,*)'
      WRITE(*,*) ' INDEXED LIST OF COVARIATES '
      WRITE(*,*) '(J,: , COVNAME(J+1),J = 1, N_VARS)'
      WRITE(*,*)

C   IF THIS IS THE FIRST MODEL BEING FITTED GOTO 22
      IF (ICOL .EQ. 0) GOTO 22
      O/W DISPLAY PREVIOUS MODELS AND THEIR DEVIANCES
      NINTS=0
      NCOV=0
      WRITE(*,*) ' PREVIOUS MODEL(S) FITTED'

      WRITE(*,*)' MDL 0: NULL MODEL'
      WRITE(*,7092) DLGLK
      WRITE(*,*)'
      DO 233 J = 1,NRUN
        WRITE(*,7091) J, (COVNAME(MODVAR(J,K)+1),K=2,MODVAR(J,1)+1)
        FORMAT('2X',MDL',1X,I2,:',1X,10(5(A1,1X),/10X))

      DO 2334 K = 1,MODINT(J,1)
        WRITE(*,(10X,A5,A5,') COVNAME(MODINT(J,2*K)+1),'*,
        + COVNAME(MODINT(J,2*K+1)+1)
        + WRITE(*,*)'
        2334 WRITE(*,7092) LIKRATIO(J)
        FORMAT(10X,'DEVIANCE',':',1X,F15.5)
        7092 WRITE(*,(' '))

      WRITE(*,8013)
      FORMAT('1X','WHICH MODEL DO YOU WISH TO USE AS A BASE? ',$)
      8013 READ(*,*)
      NSUBS=NSUBS-NUMMISS

C   RETRIEVE INFORMATION FROM BASE MODEL...E.G.: NUMBER OF COVARIATES,
C   NUMBER OF INTERACTIONS ETC.
      IF(BASE .GT. 0) THEN
        C1 = MODVAR(BASE,1)
        C2 = MODINT(BASE,1)
      ELSE
        C1=0
        C2=0
      ENDIF

      WRITE(*,8014)
      FORMAT('1X','HOW MANY NEW VARIABLES DO YOU WISH TO ADD? ',$)
      8014 GOTO 2201
      IF(BASE.GT.0) THEN
        DO 2205 K = 1,C1
          COL(K) = MODVAR(BASE,K+1)
        ENDIF

      2201 WRITE(*,7079)N_VARS
      READ(*,*) NCOV
      2205 WRITE(*,7079)
      READ(*,*) NCOV
      ENDIF

      IF(BASE.GT.0) THEN
        DO 2205 K = 1,C1
          COL(K) = MODVAR(BASE,K+1)
        ENDIF

      ELSE
        IF (NCOV .GT. 0) THEN
          WRITE(*,8015)
        ENDIF
      ENDIF

      8015 FORMAT('1X','ENTER INDEX OF CHOSEN COVARIATES: ',$)
      IF(BASE .GT. 0) THEN
        NCOV = C1 + NCOV
        IF(NCOV.GT.C1) THEN
          READ(*,*) (COL(K), K = C1+1, NCOV)
        ENDIF
      ELSE
        IF (NCOV.GT.0) THEN
          READ(*,*) (COL(K), K = 1, NCOV)
        ENDIF
      ENDIF

      MODVAR(NRUN+1,1) = NCOV
      DO 2206 K = 1,NCOV
        MODVAR(NRUN+1,K+1) = COL(K)
      2206 ICOL=ICOL+1
    
```

```

IF ((ICOL .GT. 0) .AND. (NCCV .GT. 1)) THEN
  WRITE(*,8016)
  GOTO 2202
ENDIF

FORMAT(1X 'HOW MANY NEW INTERACTIONS? (0=none) ', $)

8016  FORMAT(1X 'ENTER NUMBER OF FAIR-WISE INTERACTIONS (0=none) ', $)
      IF (NINTS.GT.0) THEN
        IF (BASE .EQ. 0) THEN
          DO 20 J = 1,NINTS
            WRITE(*,8017)
            READ(*,*) NINTS
          ENDIF
        ELSE
          MODINT(NRUN+1,1) = NINTS + C2
          DO 2207 K = 1,NINTS
            MODINT(NRUN+1,2*K) = INTERDEX(K,1)
            MODINT(NRUN+1,2*K+1) = INTERDEX(K,2)
          ELSE
            MODINT(NRUN+1,1) = MODINT(BASE,K+1)
            DO 2208 K = 1,C2
              INTERDEX(K,1) = MODINT(BASE,2*K)
              INTERDEX(K,2) = MODINT(BASE,2*K+1)
            DO 2209 K = 1,2*C2
              MODINT(NRUN+1,K+1) = MODINT(BASE,K+1)
            DO 2211 J = C2+1,C2+NINTS
              WRITE(*,8019)
              FORMAT(1X 'ENTER PAIR OF INDICES FOR ONE INTERACTION: ', $)
              READ(*,*) INTERDEX(J,1),INTERDEX(J,2)
            DO 2212 K = 1,NINTS
              MODINT(NRUN+1,2*C2+K) = INTERDEX(C2+K,1)
              MODINT(NRUN+1,2*C2+2*K) = INTERDEX(C2+K,2)
            ENDIF
          ELSE
            IF (BASE .EQ. 0) THEN
              GOTO 2214
            ENDIF
            MODINT(NRUN+1,1) = C2
            DO 2218 K = 1,C2
              INTERDEX(K,1) = MODINT(BASE,2*K)
              INTERDEX(K,2) = MODINT(BASE,2*K+1)
            DO 2219 K = 1,2*C2
              MODINT(NRUN+1,K+1) = MODINT(BASE,K+1)
              NINTS = C2
            ENDIF
          ENDIF
        ENDIF
      ENDIF
      NUMMISS=0
      DO 32 I = 1,3000
        do 32 J=1,30
          XX1(I,J)=0
      32 DO 7066 I = 1, NSUBS
        FLAG(I) = 0
      7066 DO 7083 I = 1, NSUBS
        FLAG(I) = 1, NCQV
      7083
    ENDIF
  ENDIF
  IF (XX(I,COL(J)+2).EQ.-999.00) THEN
    FLAG(I) = 1
    NUMMISS=NUMMISS+1.
  ENDIF

  C CREATE DATA MATRIX
  K=1
  DO 7067 I=1, NSUBS
    IF (FLAG(I) .EQ. 0) THEN
      XX1(K,1) = XX(I,1)
      XX1(K,2) = XX(I,2)
      DO 2 J = 1, NCQV
        XX1(K,J+2) = XX(I,COL(J)+2)
      2 ENDIF
    CONTINUE
  7067
  C CREATE INTERACTIONS
  IF (NINTS .GT. 0) THEN
    J=1
    DO 31 K = 1, NSUBS
      IF (FLAG(K) .EQ. 0) THEN
        DO 30 L = 1, NINTS
          XX1(J,L+2) = XX(K,INTERDEX(L,1)+2)
        30 +
        J=J+1
      ENDIF
    CONTINUE
  31
  C DECREMENT NUMBER OF SUBJECTS BY NUMBER WHO HAVE MISSING VALUES
  C FOR THE PARTICULAR CHOICE OF COVARIATE.
  NSUBS=NSUBS-NUMMISS
  C STORE VARIABLE NAMES
  DO 106 J = 1, NCQV
    COVRT(J) = COVNAME(COL(J)+1)
  106 DO 107 J = 1,NINTS
    IF (NINTS .GT. 0) THEN
      DO 107 J = 1,NINTS
        COVRT(J+NCQV) = COVNAME(INTERDEX(J,1)+1) (:4) // '*' //
        COVNAM(INTERDEX(J,2)+1) (:4)
      107 +
      ENDIF
    ENDIF
  C ENTER START VALUES FOR PARAMETERS OF BASELINE MODEL INTO X
  IF (BASE .EQ. 0) THEN
    DO 3123 I = 1, M - NCQV
      START(I) = BASENULL(I)
      X(I)=START(I)
    3123 ELSE
      DO 9191 I = 1, M - NCQV
        START(I) = PAMBASE(BASE,I)
        X(I)=START(I)
      9191
    ENDIF
  ENDIF
  C CREATE DATA MATRIX IGNORING THE RECORDS WITH MISSING OBSERVATIONS .
  IF (BASE .EQ. 0) THEN
    DO 3123 I = 1, M - NCQV
      START(I) = BASENULL(I)
      X(I)=START(I)
    3123 ELSE
      DO 9191 I = 1, M - NCQV
        START(I) = PAMBASE(BASE,I)
        X(I)=START(I)
      9191
    ENDIF
  ENDIF

```

C FOR REGRESSION, ENTER START VALUES FOR THE PAMS INTO X.
C THE BASELINE PARAMETERS ARE ASSIGNED THE MLE FROM THE NULL
C MODEL AS THEIR STARTING VALUES. THE REGRESSION PARAMETERS
C ARE ASSIGNED A STARTING VALUE OF ZERO.

```

C          DISPLAY ERROR MESSAGES DEPENDING ON THE VALUE OF IFAIL
C
      IF (IFAIL .EQ. 0) THEN
        WRITE(*,*) ''
        WRITE(*,6666) ITER
        WRITE(*,*)
      ELSEIF

```

```

      X(I)=0
      ELSE
        C IF (BASE .GT. 0) THEN
          C POLD = MODVAR(BASE,1) + MODINT(BASE,1)
          DO 3124 I = 1, POLD
            X(I+2)=PARMCOV(BASE,I)
          C
          POLD=0
        ENDIF
        C DO 74 I = POLD+1, NCOV
          X(I+2)=0
        ENDIF
        C
        C INITIALIZATION OF ESTIMATED S.D.s AND THE ESTIMATED COVARIANCE MATRIX
        DO 34 I = 1 , M
          ST(I)=0.0D0
          DO 33 J = 1 , M
            A(I,J)=0
            MAT(I,J) = 0.0D0
          CONTINUE
        C
        ----- SOLVE PROBLEM -----
        DO 34 I = 1 , M
          ST(I)=1
          DO 33 J = 1 , M
            A(I,J)=0
            MAT(I,J) = 0.0D0
          CONTINUE
        C
        C CALCULATE THE SCORE TEST STATISTIC
        DO 8090 I=M-NCOV+1,25
          XINC(I)=0
          MM=M
          X0(I)=0
          PARTIII(X0,A,MM,XINC,MAT,ICOV,OBJF)
          CALL SCORE(A,ICOV,XINC,STAT)
        C
        N=M
        IFAIL=-1
        COVTEMP=NCOV
        OBJF=0
        DO 7099 I=1,25
          OBGRD(I)=0
          WRITE(*,*), '-----OPTIMIZING THE LIKELIHOOD FUNCTION-----'
          WRITE(*,*), '(PLEASE WAIT)'
          CALL EO4UEP('MAJOR PRINT LEVEL = 0')
          CALL EO4UEP('MAJOR ITERATION LIMIT = 1 000')
        C
        CALL EO4UCP(N,NCLIN,NGNLN,NROWU,NROWR,AA,BL,BU,
        + EO4UDM,OBJFUN,ITER,ISTATE,C,CJAC,CLAMA,OBJF,OBJRD,R,
        + X,INWORK,LWORK,WORK,LWORK,TUSER,USER,IFAIL)
        C
        IF (IFAIL .EQ. 1) THEN
          WRITE(*,1007)
          WRITE(*,6971)
          WRITE(1007)
          GOTO 8083
        ENDIF
        C
        IFAIL=FORMAT(1X,'ERROR: THE FINAL ITERATE SATISFIES FIRST ORDER',1X,
        + 'KUHN-TUCKER CONDITIONS, BUT THE SEQUENCE OF ITERATES HAS NOT',
        + 'CONVERGED! POSSIBLE REASON: REQUESTED ACCURACY IS NOT',
        + 'ATTAINABLE WITH THE GIVEN PRECISION OF THE PROBLEM.')
        C
        IF (IFAIL .EQ. 2) THEN
          WRITE(*,1007)
          WRITE(*,6972)
          WRITE(*,1007)
        ENDIF
        C
        IFAIL=FORMAT(1X,'ERROR: NO FEASIBLE POINT EXISTS FOR THE LINEAR',1X,
        + 'CONSTRAINTS AND BOUNDS FOR THE GIVEN VALUE OF LINEAR',1X,
        + 'FEASIBILITY TOLERANCE')
        C
        IF (IFAIL .EQ. 3) THEN
          WRITE(*,1007)
          WRITE(*,6973)
          WRITE(*,1007)
          GOTO 8083
        ENDIF
        C
        IFAIL=FORMAT(1X,'ERROR: NO FEASIBLE POINT EXISTS FOR THE NON-',1X,
        + 'LINEAR CONSTRAINTS')
        C
        IF (IFAIL .EQ. 4) THEN
          WRITE(*,1007)
          WRITE(*,*), 'ITERATIONS EXCEEDED 1000!'
          WRITE(*,*), '-----'
          GOTO 8083
        ENDIF
        C
        IF (IFAIL .EQ. 6) THEN
          WRITE(*,1007)
          WRITE(*,6975)
          WRITE(*,1007)
          GOTO 8083
        ENDIF
        C
        IFAIL=FORMAT(1X,'FINAL ITERATE DOES NOT SATISFY FIRST',1X,
        + 'ORDER KUHN-TUCKER CONDITIONS')
        C
        IF (IFAIL .EQ. 7) THEN
          WRITE(*,1007)
          WRITE(*,6976)
        ENDIF
      ENDIF
    ENDIF
  ENDIF
ENDIF

```

```

        WRITE(*,1007)
        GOTO 8083
      ENDIF

6976   + FORMAT('1X, 'USER-PROVIDED DERIVATIVES OF THE',1X,
      + 'OBJECTIVE FUNCTION APPEAR TO BE INCORRECT')
      IF (IAFAIL .EQ. 9) THEN
        WRITE(*,1007)
        WRITE(*,6977)
        WRITE(*,1007)
        GOTO 8083
      ENDIF

6977   FORMAT(' AN INPUT PARAMETER IS INVALID')

      C NEGATE THE OBJECTIVE FUNCTION ON SUCCESSFUL CONVERGENCE
      C SINCE THE NAG ROUTINE DOES A MINIMIZATION
      OBJF=-OBJF
      GOTO 8084

      C IF CONVERGENCE HAS BEEN REACHED FOR THE REG. MODEL , EVALUATE L.R.
      8083  WRITE(*,8020)
      8020  FORMAT(' DO YOU WISH TO CHANGE THE MODEL? (Y/N) ',\$)
            READ(*,'(A)')  ANS
            IF ((ANS.EQ.'Y').OR.(ANS.EQ.'Y').OR.(ANS.EQ.'YES')).OR.
            +(ANS.EQ.'yes') THEN
              ICOI=ICOL-1
              DO 8082 I=1,25
                MODVAR(NRUN+1,I)=0
                MODINT(NRUN+1,I)=0
                GOTO 7081
              ENDIF
              IF ((ANS.EQ.'N').OR.(ANS.EQ.'n').OR.(ANS.EQ.'NO')).OR.
              +(ANS.EQ.'no') THEN
                WRITE(*,*), EXITING THE PROGRAM'
                GOTO 9
              ELSE
                WRITE(*,*), ' INVALID ANSWER!'
                GOTO 8083
              ENDIF

8084  CHI = DLGLK - (-2*OBJF)
      LIKRATIO(NRUN+1)=-2*OBJF

      C DECLARE SUCCESSFUL CONVERGENCE AND DISPLAY NUMBER OF ITERATIONS NEEDED TO
      C REACH CONVERGENCE AND THE VALUE OF THE OBJECTIVE FUNCTION

3115  DO 3115 K = 1,2
      PARMBASE(NRUN+1,K) = X(K)
      DO 3116 K = 3,M
      PARMCOV(NRUN+1,K-2) = X(K)

      NRUN = NRUN + 1

      IF (NRUN .GT. 25) THEN
        WRITE(*,*),'MAXIMUM NUMBER OF MODEL ATTEMPTS HAS BEEN',
        +'REACHED.'
        WRITE(*,*),'START THE PROGRAM OVER'
        GOTO 9
      ENDIF

      C COMPUTE SAMPLE INFORMATION MATRIX
      CALL PARTII(X,A,M,XINC,MAT,NCOV,OBJF)

      C INVERT THE SAMPLE INFORMATION MATRIX (-A = MAT).
      CALL ARINV ( M,MAT )

      C OUTPUT: SPECIFY WHICH MODEL HAS BEEN USED IN THE ANALYSIS.
      WRITE(25,1007)
      WRITE(25,*), 'MLE FROM MIXED POISSON REGRESSION MODEL'
      WRITE(25,*)

      C WRITE THE NAMES OF COVARIATES ENTERING ANALYSIS.
      WRITE(25,1014) (COVRT(K),K = 1 , NCOV)
      WRITE(25,*)
      WRITE(25,*)

      C EVALUATE THE ESTIMATED S.D. OF THE M.L.ES OF THE PARMS.
      DO 81 I = 1,M
        IF(MAT(I,I) .LE. 0.0) THEN
          WRITE(*,*), ' INFORMATION MATRIX IS SINGULAR'
          GO TO 8083
        ENDIF
      81   ST(I)=SQRT(MAT(I,I))

      C WRITE TO THE OUTPUT FILE THE MLE AND SDS OF THE PARMS OF THE MIXING DISTN
      WRITE(25,8005) NSUBS
      WRITE(25,*)
      WRITE(25,*)

      WRITE(*,8005) NSUBS
      WRITE(*,*)
      WRITE(25,*)

      WRITE(25,8006)
      WRITE(25,*)
      WRITE(*,8006)
      WRITE(*,*)

      WRITE(25,1015)X(1),X(2),ST(1),ST(2)
      WRITE(25,*)
      WRITE(*,1015)X(1),X(2),ST(1),ST(2)
      WRITE(*,*)

      C IF THE MODEL CONTAINS COVARIATES THEN WRITE TO FILE THE VALUES OF THE
      C L.R., SCORE STATISTIC AND D.F.
      IF (NCOV .GT. 0 ) THEN
        WRITE(25,1018) -2*OBJF,DGLK
        WRITE(*,1018) -2*OBJF,DGLK
        WRITE(25,1019) CHI,ICOV
        WRITE(*,1019) CHI,ICOV

```

```

      WRITE(25,1020) STAT, ICOV
      WRITE(*,1020) STAT, ICOV
      WRITE(25,1021)
      WRITE(*,1021)

C EVALUATE THE Z-VALUE FOR EACH COVARIATE (MLE/S.E.).
C EVALUATE THE NORMAL-BASED 95% C.I. FOR EACH REGRESSION COEFFICIENT
C IF THE 95% C.I. EXCLUDES ZERO THEN FLAG THE COVARIATE AS SIGNIFICANT.

      IFLAG=0

      DO 90 J = 3 , M
      D2=0.0
      C1L=0.0
      C1U=0.0
      IF(ABS(DZ).GT. 1.96) THEN
      IFLAG=1
      WRITE(25,1022) COVRT(J-2), X(J), ST(J), DZ, C1L, C1U
      WRITE(*,1022) COVRT(J-2), X(J), ST(J), DZ, C1L, C1U
      ELSE
      WRITE(25,1023) COVRT(J-2), X(J), ST(J), DZ, C1L, C1U
      WRITE(*,1023) COVRT(J-2), X(J), ST(J), DZ, C1L, C1U
      ENDIF
      CONTINUE
      90
      CONTINUE
      10
      CONTINUE

      IF(IFLAG.EQ. 1) THEN
      WRITE(25,*)
      WRITE(25,'*')
      * : SIGNIFICANT AT THE 5% LEVEL'
      WRITE(*,*)
      WRITE(*,*)
      * : SIGNIFICANT AT THE 5% LEVEL'
      ENDIF
      WRITE(25,1024)
      ELSE
      WRITE(*,1025) -OBJF
      ENDIF
      1025 FORMAT(6X,'==== -LOG-LIKELIHOOD =',F12.5 )
      C
      WRITE TWO OUTPUT THE CORRELATION MATRIX.
      WRITE(25,'(23X,30H ESTIMATED CORRELATION MATRIX:) ')
      WRITE(25,*)
      WRITE(25,*)
      WRITE(25,'(7X,30I7 ) (K, K=1,M)
DO 110 I = 1 , M-NCOV
      WRITE(25,'(5X,A5,30(1X,F6.3 ))') GMA(I),
      + ( MAT(I,J)/(ST(I)*ST(J))),J=1,M)
      DO 12 I = M-NCOV+1 , M
      WRITE(25,(A10.30(1X,F6.3 ))') COVRT(I-M+NCOV),
      + ( MAT(I,J)/(ST(I)*ST(J))),J=1,M)
      IF(IFLG.EQ.1) THEN
      GOTO 7078
      ENDIF
      CONTINUE
      GO TO 7078

      99
      *****
      FORMAT STATEMENTS *****
      C
      *****
      FORMAT(2X,'STARTING VALUES FOR NULL MODEL.', ('1X,I2,1X,'PAR.)?' )
      1005 FORMAT(2X,'STARTING VALUES FOR CURRENT MODEL, ('16,1X,
      + 'PARAMETERS) ? ' )
      1006 FORMAT(2X,'STARTING VALUES FOR CURRENT MODEL, ('16,1X,
      + 'PARAMETERS) ? ' )

      1007 FORMAT(39H -----
      + 39H -----, GRADIENT VECTOR: ')
      1009 FORMAT('=====,5X,'
      1010 FORMAT('=====,7H ITER =,I4,6X,'-OBJF =',F16.5,7X,8H COND =
      + E16.5,/, '=====
      1011 FORMAT(32H =====,XO(1),...,XO(12,1H) /,10(8F15.8)
      1012 FORMAT(25H SUCCESSFUL CONVERGENCE: 10X,'OBJF =',F12.4 )
      1013 FORMAT('2X,'NEW VALUES .('1X,I2,1X,'PAR.)? ')
      1014 FORMAT(' COVARIATES ENTERING ANALYSIS: ',4A12,/,30(3OX,4A12,/
      + ' .)
      1015 FORMAT(16X,'SHAPE =',F12.5,1OX,' MEAN =',F11.6,/,16X,'ST.DEV =
      + (' , F11.5,/, '10X,'ST.DEV =',(' , F10.6,'))
      1016 FORMAT(7X,'SHAPE =',F9.4,8X,' MEAN =',F11.6,5X,1ST-GAP=
      + ,F9.1,/,7X,'ST.DEV =',(' , F9.5,/) ,6X,'ST.DEV =',(' , F10.6
      + ') ,5X,'ST.DEV =',(' , F9.5,/')
      1017 FORMAT(6X,'DEVIANCE FOR NULL MODEL (-2*(LOG-LIK)) = ',F13.5)
      1018 FORMAT(6X,'DEVIANCE: CURRENT MODEL',3X,' =',F13.5,' WITH ',I3,
      + F13.5)
      1019 FORMAT(6X,'LIK. RATIO TEST STATISTIC',2X,' =',F13.5,' WITH ',I3,
      + ' DEGREE(S) OF FREEDOM')
      1020 FORMAT(6X,'SCORE TEST STATISTIC',7X,' = ',F12.5,' WITH ',I3,
      + ' DEGREE(S) OF FREEDOM')
      1021 FORMAT(6X,'COVARIATE',4X,' M.L.E.',4X,' ST.DEV.',5X,
      + 'Z-VALUE',4X,' 95% C. INTERVAL',6X,'-----',4X,'-----',
      + 4X,'-----',5X,'-----',4X,'-----')
      1022 FORMAT(4X,A10.5X,F9.6,'*',3X,F9.6,4X,F7.4,4X,'( ',F7.3,', ',
      + ')')
      1023 FORMAT(4X,A10.5X,F9.6,4X,F9.6,4X,F7.4,4X,'( ',F7.3,', ',
      + ')')
      1024 FORMAT(5X,58(' -'))
      7079 FORMAT(' HOW MANY COVARIATES DO YOU WISH TO USE? (1 --',
      + I3, ') ',\$)
      9 STOP
      END

      SUBROUTINE OBJFUN(MODE,N,X,OBJF,OBJGRD,NSTATE,IUSER,USER)
      C THIS SUBROUTINE COMPUTES THE OBJECTIVE FUNCTION, AND THE FIRST
      C PARTIAL DERIVATIVE FOR THE NAG SUBROUTINE
      INTEGER MODE, N, NSTATE, IUSER(1), USER(1)
      DOUBLE PRECISION X(25), OBJF, OBJGRD(25)
      DOUBLE PRECISION MAT(25,25), XINC(25), A(25,25)
      INTEGER K, COV
      COMMON XX(3000,30), XX1(3000,30), P(25,25),
      + MDL, NSUBS, START(5), COVTEMP
      DO 669 K=1,25
      XINC(K)=0
      DO 671 K=1,N
      OBJF=-OBJF
      CALL PARTII(X,A,N,XINC,MAT,COV,OBJF)
      RETURN
      END

      671 DO 672 OBJGRD(K) = XINC(K)
      672 RETURN
      END

```

```

SUBROUTINE OBJNULL(MODE,N,XO,OBJF,OBJGRD,NSTATE,TUSER,USER)
C THIS SUBROUTINE COMPUTES THE OBJECTIVE FUNCTION, AND THE FIRST
C PARTIAL DERIVATIVE FOR THE NAG SUBROUTINE
COMMON +
      INTEGER MODE,N, NSTATE,USER(1),TUSER(1)
      DOUBLE PRECISION XO(25), OBJF,OBJGRD(25)
      DOUBLE PRECISION MAT(25,25),XINC(25),A(25,25)
      INTEGER K,COV
      XX(3000,30),XX1(3000,30),P(25,25),
      MDL, NSUBS, START(5), COVTEMP

DO 666 K=1,25
  XINC(K)=0
666   COV=COVTEMP
      CALL PARTIII(XO,A,N,XINC,MAT,COV,OBJF)
      OBJF=-OBJF
      DO 668 K=1,N
        OBJGRD(K) = XINC(K)
668   RETURN
670   END

```

```

SUBROUTINE DECOMP(MDIM,M,A,COND,IPVT,WORK)
  INTEGER MDIM,M,IPVT(MDIM)
  REAL*8 COND,WORK(MDIM),A(25,25)
  REAL*8 EK,T,ANORM,YNORM,ZNORM
  INTEGER MM1,I,J,K,KP1,KB,KM1,M1

      IPVT(M) = 1
      IF ( M .EQ. 1 ) GO TO 80
      MM1 = M-1
      ANORM = 0.0D0
      DO 10 J = 1 , M
        T = 0.0D0
        DO 5 I = 1 , M
          T = T + DABS( A(I,J) )
        CONTINUE
        IF ( T .GT. ANORM ) ANORM = T
      5
      DO 35 K = 1 , MM1
        KP1 = K + 1
        M1 = K
        DO 15 I = KP1 , M
          IF (DABS(A(I,K)) .GT. DABS(A(M1,K)) ) M1 = I
        CONTINUE
        IPVT(K) = M1
        IF ( M1 .NE. K ) IPVT(M) = -IPVT(M)
        T = A(M1,K)
        A(M1,K) = A(K,K)
        A(K,K) = T
        IF ( T .EQ. 0.0D0 ) GO TO 35
        DO 20 I = KP1 , M
          A(I,K) = -A(I,K) / T
        CONTINUE
        DO 30 J = KP1 , M
          T = A(M1,J)
          A(M1,J) = A(K,J)
          A(K,J) = T
        IF ( T .EQ. 0.0D0 ) GO TO 30
        DO 25 I = KP1 , M
          A(I,J) = A(I,J) + A(I,K) * T
        CONTINUE
        CONTINUE
        CONTINUE
      20
      25
      30
      35
      DO 50 K = 1 , M
        EK = 1.0D0
        IF ( T .LT. 0.0D0 ) EK = -1.0D0
        KM1 = K - 1
        DO 40 I = 1 , KM1
          T = T + A(I,K) * WORK(I)
        CONTINUE
        DO 60 KB = 1 , MM1
          K = M - KB
          T = 0.0D0
          KP1 = K + 1
          DO 55 I = KP1 , M
            T = T + A(I,K) * WORK(K)
          CONTINUE
          WORK(K) = T
        40
        45
        50
        55
        M1 = IPVT(K)

```

Thu Jun 9 13:42:25 1994

```

C ON OUTPUT, THE SUBROUTINE RETURNS THE VALUE OF THE SCORE TEST
C C STATISTIC "STAT".
C
C REAL*8 A(25,25),BTMB(25,25),MAT(25,25),BTM(25,25),GOFF(25),
C + STAT,U(25),XINC(25)
C
C INTEGER MC,ICOV
C COMMON XX(3000,30),XX1(3000,30),P(25,25),
C + MDL,NSUBS,START(5),COVTEMP
C
C MC = 2
C
C CALL SOLVE ( MDIM,M,A,WORK,IPVT )
C
C ZNORM = 0.0D0
C DO 70 I = 1 , M
C   ZNORM = ZNORM + DABS( WORK(I) )
C CONTINUE
C COND = ANORM*ZNORM/YNORM
C IF (COND .LT. 1.0D0 ) COND = 1.0D0
C RETURN
C COND = 1.0D0
C IF ( A(1,1) .NE. 0.0D0 ) RETURN
C COND = 1.0E+3.2
C RETURN
C END
C
C SUBROUTINE SCOLVE(MDIM,M,A,XINC,IPVT)
C INTEGER MDIM,M,IPVT(MDIM)
C REAL*8 A(MDIM,MDIM),XINC(MDIM)
C INTEGER KB,KM1,MM1,PK1,I,K,M1
C REAL*8 T
C
C IF ( M .EQ. 1 ) GO TO 50
C MM1 = M - 1
C DO 20 K = 1 , MM1
C   PK1 = K + 1
C   M1 = IPVT(K)
C   T = XINC(M1)
C   XINC(M1) = XINC(K)
C   XINC(K) = T
C   DO 10 I = KPK1 , M
C     XINC(I) = XINC(I) + A(I,K)*T
C   CONTINUE
C   DO 40 KB = 1 , MM1
C     KM1 = M - KB
C     K = KM1 + 1
C     XINC(K) = XINC(K)/A(K,K)
C     T = -XINC(K)
C     DO 30 I = 1 , KM1
C       XINC(I) = XINC(I) + A(I,K)*T
C     CONTINUE
C     XINC(1) = XINC(1)/A(1,1)
C   RETURN
C   END
C
C SUBROUTINE SCORE(A,ICOV,XINC,STAT)
C THIS SUBROUTINE EVALUATES THE SCORE TEST STATISTIC FOR TESTING
C THE NULL HYPOTHESIS THAT THE REGRESSION COEFFICIENTS IN THE
C REGRESSION MODEL ARE ALL ZEROS. THE FOLLOWING QUANTITIES MUST
C BE PASSED AS INPUT:
C
C 1. SAMPLE INFORMATION MATRIX = -A(25,25).
C 2. SCORE VECTOR AT THE MLE VALUES = -XINC(25,25).
C 3. NUMBER OF COVARIATES IN THE MODEL = ICOV.
C 4. THE NUMBER OF BASELINE PARAMETERS. = MC.
C
C C ON OUTPUT, THE SUBROUTINE RETURNS THE VALUE OF THE SCORE TEST
C C STATISTIC "STAT".
C
C DO 10 I = 1 , 25
C   U(I) = -XINC(I)
C DO 1 J = 1 , 25
C   BTM(I,J) = 0.0D0
C   BTMB(I,J) = 0.0D0
C   MAT(I,J) = 0.0D0
C   A(I,J) = -A(I,J)
C CONTINUE
C
C STEP 1: FIND THE COVARIANCE MATRIX FOR THE BASELINE MODEL
C
C DO 2 J = 1 , MC
C   DO 2 I = 1 , MC
C     MAT(I,J) = A(I,J)
C   CONTINUE
C   CALL ARINV(MC,MAT)
C
C STEP 2: EVALUATE TRANS(TGB)*MAT*TGB
C
C DO 3 I = 1 , ICOV
C   DO 3 J = 1 , MC
C     DO 3 K = 1 , MC
C       BTM(I,J) = BTM(I,J) + A(I+MC,K)*MAT(K,J)
C     CONTINUE
C     CALL ARINV(MC,MAT)
C   CONTINUE
C   DO 4 I = 1 , ICOV
C     DO 4 J = 1 , ICOV
C       DO 4 K = 1 , MC
C         BTMB(I,J) = BTMB(I,J) + BTM(I,J) + BTM(I,K)*A(K,J+MC)
C     CONTINUE
C   CONTINUE
C
C STEP 3: EVALUATE PARTITIONED INVERSE
C
C DO 5 I = 1 , ICOV
C   GOFF(I) = 0.0
C   DO 5 J = 1 , ICOV
C     MAT(I,J) = A(I+MC,J+MC) - BTMB(I,J)
C   CONTINUE
C   CALL ARINV ( ICOV,MAT )
C
C STEP 4: CALCULATE THE SCORE TEST STATISTIC
C
C STAT = 0.0
C DO 6 J = 1 , ICOV
C   DO 6 I = 1 , ICOV
C     GOFF(J) = GOFF(J) + U(I+MC)*MAT(I,J)
C   CONTINUE
C   DO 7 I = 1 , ICOV
C     STAT = STAT + GOFF(I)*U(I+MC)
C   CONTINUE
C
C RETURN
C END
C
C
```

Thu Jun 9 13:42:25 1994

```

C SUBROUTINE USPKD (PACKED,NCHARS,UNPAKD,NCHMTB)
C          SPECIFICATIONS FOR ARGUMENTS
C
C      INTEGER      NC, NCHARS, NCHMTB
C      INTEGER      UNPAKD(nchars), PACKED(nchars), LBYTE, LBLANK
C      INTEGER*2     EQUIVALENCE (LBYTE, IBYTE)
C      DATA         IBLANK /1H /
C      DATA         IBYTE /1H /
C      DATA         IBLANK /1H /
C
C      NCHMTB = 0
C
C      IF (NCHARS.LE.0) RETURN
C
C      NC = 129
C      IF (NCHARS.LT.NC) NC=NCHARS
C      NWORDS = NC+NC+NC-NC
C      J = 1
C      DO 110 I = 1,NWORDS,4
C          UNPAKD(I) = PACKED(J)
C          UNPAKD(I+1) = LBLANK
C          UNPAKD(I+2) = LBLANK
C          UNPAKD(I+3) = LBLANK
C 110  J = J+1
C
C      CHECK UNPAKD ARRAY AND SET NCHMTB BASED ON TRAILING BLANKS FOUND
C
C      DO 200 N = 1,NWORDS,4
C          NN = NWORDS - N - 2
C          LBYTE = UNPAKD(NN)
C          IF (IBYTE .NE. IBLANK) GO TO 210
C 200  CONTINUE
C
C      NN = 0
C 210  NCHMTB = (NN + 3) / 4
C      RETURN
C      END
C
C      SUBROUTINE PARTII (XO,A,M,XINC,MAT,COV,OBJF)
C
C      SCORE VECTOR AND SAMPLE INFORMATION MATRIX CALCULATIONS FOR :
C          PARETO MODEL, MODEL-I AND MODEL-II
C
C      THIS SUBROUTINE EVALUATES THE FIRST AND SECOND PARTIAL
C      DERIVATIVES OF THE LOG LIKELIHOOD FUNCTION FOR THE PARETO
C      MODEL, MODEL-I AND MODEL-II
C
C      THE VARIABLES IN THIS SUBROUTINE ARE DEFINED AS FOLLOWS:
C
C      XO(25) = VECTOR OF THE UPDATED PARAMETER ESTIMATES WITH:
C          A(25,25) = P(25,25) = MATRIX OF SECOND PARTIAL
C          DERIVATIVES OF THE LOG LIKELIHOOD FUNCTION.
C
C      XO(1) = MLE OF SHAPE PAR. OF THE GAMMA MIXING DISTRIBUTION.
C      XO(2) = MLE OF MEAN OF THE GAMMA MIXING DISTRIBUTION.
C      XO(3) = MLE OF THE 1\030TH.022 REGRESSION COEFFICIENT IN THE
C          REGRESSION MODEL.
C
C      XINC(25) = NEGATIVE OF THE SCORE VECTOR
C          A(25,25) = P(25,25) = MATRIX OF SECOND PARTIAL
C          DERIVATIVES OF THE LOG LIKELIHOOD FUNCTION.
C
C      OBJF = THE OBJECTIVE FUNCTION TO BE MAXIMIZED
C          = LOG LIKELIHOOD FUNCTION.
C
C      MDL = A SCALAR INDICATOR OF MODEL TO BE USED IN ANALYSIS WITH
C          3=MODEL-II.
C          1=PARETO MODEL
C          VALUES:
C          START = VECTOR OF STARTING VALUES FOR THE BASELINE PARAMETERS.
C
C      Gi = T(i,2) + T(i,3) - 1 = NUMBER OF TUMOR VISITS FOR THE
C          iTH PATIENT.
C      XB1 = Zi * Beta
C      XB = EXP(XB1)
C      SMI = T(i,1) + T(i,5) = TOTAL FOLLOW UP TIME.
C      DEN = XO(2)*XB*SMI + XO(1)
C      GB = Gi + XO(1)
C
C      REAL*8      SMI, SM2, SM3, DEN, GI, OBJF, XB, XB1, MAT(25,25),
C                  SMI_A(25,25), XO(25), XINC(25), GB
C
C      INTEGER      COV, S, M
C      COMMON      XX(3000,30), XX1(3000,30), P(25,25),
C                  MDL, NSUBS, START(5), COVTEMP
C
C      CHECK IF THE UPDATED ESTIMATE OF THE SHAPE OR SCALE PARAMETERS IS
C      ZERO, IF SO THEN REPLACE THE UPDATED VALUE BY THE STARTING VALUE.
C
C      IF (XO(1).LE. 0.0) XO(1) = START(1)
C      IF (XO(2).LE. 0.0) XO(2) = START(2)
C
C      INITIALIZATION OF THE SCORE VECTOR AND THE MATRIX OF SECOND PARTIAL
C      DERIVATIVES OF THE LOG LIKELIHOOD FUNCTION.
C
C      OBJF = 0.0DO
C      DO 4 I = 1 , 25
C          XINC(I) = 0.0DO
C          DO 4 J = 1 , 25
C              P(I,J) = 0.0DO
C          CONTINUE
C 4
C
C      DO 10 I = 1 , NSUBS
C
C      C      INITIALIZE THE DIFFERENT SCALARS NEEDED IN CALCULATIONS .
C
C      SM1=0.0DO
C      SM2=0.0DO
C      SM3=0.0DO
C      DEN=0.0DO
C      GI = 0.0DO
C      GB = 0.0DO
C      SMI= 0.0DO
C      XB =1.0DO
C      XB1=0.0DO
C
C      SET UP THE VALUES OF Gi AND SMI . THE VALUES OF THESE
C      QUANTITIES DEPEND ON THE MODEL CHOSEN .
C
C      GI = XX1(I,1)
C      SMI = XX1(I,2)
C
C      IF THE NUMBER OF FAILURES FOR THE ith PATIENT IS > 0 THEN
C          COMPUTE SMI , SM2 AND SM3 .
C
C      IF ( GI.GT. 0 ) THEN
C          DO 2 J = 1 , GI
C              SM1 = SM1 + DLOG( XO(1)+J-1 )
C              SM2 = SM2 + 1/(XO(1)+J-1)
C              SM3 = SM3 + 1/((XO(1)+J-1)*2)
C
C      END
C
```

```

2      CONTINUE
ENDIF

C COMPUTE THE SCALAR XB1, XB, DEN, GB, SCORE VECTOR AND THE
C NEGATIVE OF THE SAMPLE INFORMATION MATRIX.
C
DO 3 J = 1 , COV
  XB1 = XB1 + XX1(I,J+2)*XO(J+2)
CONTINUE
XB=DEXP(XB1)

C DEN = XB*SMI*XO(2) + XO(1)
GB = XO(1) + GI
CONTINUE

C XINC(1)=XINC(1) - ( SM2 - DLOG(DEN) +
  + (1+DLOG(XO(1)) - GB/DEN) )
C XINC(2)=XINC(2) - ( GI/XO(2) - GB*XB*SMI/DEN )
CONTINUE

DO 5 J = 1 , COV
  XINC(J+2)=XINC(J+2) - { GI*XX1(I,J+2) -
  GB*XO(2)*XX1(I,J+2)*XB*SMI/DEN
  + CONTINUE
C P(1,1)=P(1,1) - SM3 + ( 1/XO(1) - (2*DEN - GB) / (DEN**2) )
  P(1,2)=P(1,2) - XB*SMI/DEN + GB*SMI*XB / (DEN**2)
DO 7 J = 1 , COV
  P(1,J+2)=P(1,J+2) - XO(2)*XB*SMI*XX1(I,J+2)/DEN +
  GB*XO(2)*SMI*XX1(I,J+2)*XB / (DEN**2)
  + CONTINUE
C P(2,2) = P(2,2) - GI / (XO(2)**2) + GB* ( XB*SMI / DEN ) **2
DO 8 J = 1 , COV
  P(2,J+2)=P(2,J+2) - XO(1) *GB*XX1(I,J+2)*XB*SMI / (DEN**2)
  + CONTINUE
C DO 9 S = 1 , COV
  DO 9 K = S , COV
    P(S+2,K+2)=P(S+2,K+2) - XO(1) *XO(2) *GB*XX1(I,S+2)
    + CONTINUE
C OBJF = OBJF + GI* ( DLOG(XO(2)) + XB1 ) + SMI -
  + GB*DLOG( DEN ) + XO(1)*DLOG(XO(1))
  + CONTINUE
C DO 11 J = 1 , M
  DO 11 I = J , M
    P(I,J) = P(J,I)
  CONTINUE
C DO 12 I = 1 , M
  DO 12 J = 1 , M
    A(I,J) = P(I,J)
    MAT(I,J) = -P(I,J)
  CONTINUE
C RETURN
END

SUBROUTINE ARINV ( M, MAT )
C INTEGER IJOB,N,IA,IER
REAL*8 MAT(M,M),D1,D2,WKAREA(30),B(1)
N = M
IA = 25
IJOB = 1
D1 = -1.0
CALL LINV3F( MAT,B,IJOB,N,IA,D1,D2,WKAREA,IER )

SUBROUTINE LINV3F ( A,B,IJOB,N,IA,D1,D2,WKAREA,IER)
C DOUBLE PRECISION A(IA,1),B(1),WKAREA(1),C1,C2,D1,D2,WA,ZERO,
* ONE,SUM,C DATA ZERO/0.0D0/,ONE/1.0D0/
C CALL LUDETN ( A,IA,N,A,IA,0,C1,C2,WKAREA,WKAREA,WA,IER )
C CALL LUDETN ( A,IA,N,A,IA,0,C1,C2,WKAREA,WKAREA,WA,IER )
IF (D1 .LT. ZERO .AND. IJOB .GE. 1 .AND. IJOB .LT. 4) GO TO 5
D1 = C1
D2 = C2
C 5 IF (IER .GE. 128) GO TO 60
IF (IJOB .LE. 0 .OR. IJOB .GT. 4) GO TO 55
SOLVE AX = B
C IF (IJOB .EQ. 2 .OR. IJOB .EQ. 3) CALL LUJLMN ( A, IA, N, B, WKAREA, B )
IF (IJOB .NE. 1 .AND. IJOB .NE. 3) GO TO 9005
C A(N,N) = ONE/A(N,N)
NM1 = N-1
IP (NM1 .LT. 1) GO TO 9005
DO 40 II=1,NM1
L = N-II
M = L+1
DO 15 I=M,N
SUM = ZERO
DO 10 K=M,N
SUM = SUM-A(I,K)*A(K,L)
  + CONTINUE
DO 20 I=M,N
SUM = SUM-A(I,L)
  + CONTINUE
A(I,L) = WKAREA(N+I)
  + CONTINUE
DO 30 J=L,N
SUM = ZERO
IF (J .EQ. L) SUM = ONE
DO 25 K=M,N
SUM = SUM-A(L,K)*A(K,J)
  + CONTINUE
IF (Wkarea(N+J) = SUM/A(L,L))
  + CONTINUE
DO 35 J=L,N
A(L,J) = WKAREA(N+J)
  + CONTINUE
A(L,L) = WKAREA(N+J)
  + CONTINUE
DO 40 I=1,N
J = N-I+1
JP = WKAREA(J)
IF (J .EQ. JP) GO TO 50
DO 45 K=1,N
C = A(K,JP)
C PERMUTE COLUMNS OF A INVERSE
C DO 50 I=1,N
J = N-I+1
JP = WKAREA(J)
IF (J .EQ. JP) GO TO 50
DO 45 K=1,N
C = A(K,JP)

```

```

      A(K,JP) = A(K,J)
      A(K,J) = C
      CONTINUE
      45   CONTINUE
      50   CONTINUE
      GO TO 9005
      55   CONTINUE
      WARNING WITH FIX - LUJOB WAS SET INCORRECTLY
      GO TO 9000
      TERMINAL ERROR - MATRIX A IS ALGORITHMICALLY SINGULAR
      IER = 65
      GO TO 130
      60   IER = 130
      CONTINUE
      CALL UERTST(IER, 6HINV3F)
      9005 RETURN
      END

      SUBROUTINE LUDATN (A,IA,N,LU,ILU,LDGT,D1,D2,APVT,EQUIL,WA,IER)
      DIMENSION A(IA,1),LU(ILU,1),APVT(1),EQUIL(1)
      DOUBLE PRECISION A,LU,D1,D2,EQUIL,WA,ZERO,ONE,FOUR,SIXTN,SIXTH,
      *          RN,WREL,BIGA,BIG,P,SUM,AI,WI,T,TEST,Q,APVT
      *          DATA ZERO,ONE,FOUR,SIXTN,SIXTH/0.0D+1.0D+4.0D,
      *          16.0D+0.0625D0 /
      C          FIRST EXECUTABLE STATEMENT
      C          INITIALIZATION
      TER = 0
      RN = N
      WREL = ZERO
      D1 = ONE
      D2 = ZERO
      BIGA = ZERO
      DO 10 I=1,N
      BIG = ZERO
      DO 5 J=1,N
      P = A(I,J)
      LU(I,J) = P
      P = DABS(P)
      IF (P .GT. BIG) BIG = P
      5   CONTINUE
      IF (BIG .GT. BIGA) BIGA = BIG
      IF (BIG .EQ. ZERO) GO TO 110
      EQUIL(I) = ONE/BIG
      10  CONTINUE
      DO 105 J=1,N
      J1 = J-1
      IF (J1 .LT. 1) GO TO 40
      COMPUTE U(I,J), I=1,...,J-1
      105 CONTINUE
      SUM = DABS(SUM)
      WI = ZERO
      IF (IM1 .LT. 1) GO TO 20
      IM1 = I-1
      IF (IDGT .EQ. 0) GO TO 25
      COMPUTE U(I,J) WITH ACCURACY TEST
      AI = DABS(SUM)
      WI = WI+DABS(T)
      SUM = SUM-T
      WI = WI+DABS(T)
      15   CONTINUE
      LU(I,J) = SUM
      IF (AI .EQ. ZERO) AI = BIGA
      TEST = WI/AI
      WREL = TEST
      GO TO 35
      20   AI = J+1
      IF (UP1 .GT. N) GO TO 105
      DIVIDE BY PIVOT ELEMENT U(J,J)
      P = LU(J,J)
      DO 100 I=J+1,N
      LU(I,J) = LU(I,J)/P
      GO TO 105
      CONTINUE
      35   CONTINUE
      40   P = ZERO
      COMPUTE U(J,J) AND L(I,J), I=J+1,...,
      DO 70 I=J,N
      SUM = LU(I,J)
      IF (IDGT .EQ. 0) GO TO 55
      IF (IDGT .EQ. 0) WITH ACCURACY TEST
      AI = DABS(SUM)
      WI = ZERO
      IF (JM1 .LT. 1) GO TO 50
      DO 45 K=1,JM1
      T = LU(I,K)*LU(K,J)
      SUM = SUM-T
      WI = WI+DABS(T)
      CONTINUE
      LU(I,J) = SUM
      WI = ZERO
      IF (JM1 .LT. 1) GO TO 50
      DO 45 K=1,JM1
      T = LU(I,K)*LU(K,J)
      SUM = SUM-T
      WI = WI+DABS(T)
      CONTINUE
      LU(I,J) = SUM
      WI = WI+DABS(SUM)
      SUM = SUM-LU(I,K)*LU(K,J)
      CONTINUE
      LU(I,J) = SUM
      WI = WI+DABS(SUM)
      SUM = SUM-LU(I,K)*LU(K,J)
      CONTINUE
      LU(I,J) = SUM
      WI = WI+DABS(SUM)
      SUM = SUM-LU(I,K)*LU(K,J)
      CONTINUE
      LU(I,J) = SUM
      Q = EQUIL(I)*DABS(SUM)
      IF (P .GE. Q) GO TO 70
      P = Q
      IMAX = I
      DO 75 K=1,N
      P = LU(IMAX,K)
      LU(I,J,K) = P
      CONTINUE
      EQUIL(IMAX) = EQUIL(J)
      APVT(J) = IMAX
      D1 = D1*LU(J,J)
      D1 = D1*SIXTN
      IF (DABS(D1) .LE. ONE) GO TO 90
      D1 = D1*SIXTH
      D2 = D2+FOUR
      GO TO 85
      80   IF (DABS(D1) .GE. SIXTH) GO TO 95
      D1 = D1*SIXTN
      D2 = D2-FOUR
      GO TO 90
      85   IF (DABS(D1) .LE. ONE) GO TO 90
      D1 = D1*SIXTH
      D2 = D2+FOUR
      GO TO 85
      88   IF (APVT(J) .EQ. RN) GO TO 90
      D1 = D1*LU(J,K)
      LU(I,J,K) = P
      CONTINUE
      EQUIL(IMAX) = EQUIL(J)
      APVT(J) = IMAX
      D1 = D1*LU(J,J)
      D1 = D1*SIXTN
      IF (DABS(D1) .LE. ONE) GO TO 90
      D1 = D1*SIXTH
      D2 = D2-FOUR
      GO TO 95
      90   IF (DABS(D1) .GE. SIXTH) GO TO 95
      D1 = D1*SIXTN
      D2 = D2-FOUR
      GO TO 90
      95   IF (UP1 .GT. N) GO TO 105
      DIVIDE BY PIVOT ELEMENT U(J,J)
      P = LU(J,J)
      DO 100 I=J+1,N
      LU(I,J) = LU(I,J)/P
      GO TO 105
      CONTINUE
      100  CONTINUE
      WITHOUT ACCURACY

```

```

C   PERFORM ACCURACY TEST
C     IF (IDGT .EQ. 0) GO TO 9005
P = 3*N+3
WA = P*WREL
IF (WA .GT. 0.0 * (-IDGT)) .NE. WA) GO TO 9005
GO TO 9000
C   ALGORITHMIC SINGULARITY
110 IER = 129
D1 = ZERO
D2 = ZERO
9000 CONTINUE
C   CALL UERTST(IER,6HLDATN)          PRINT ERROR
9005 RETURN
C   SUBROUTINE LUELBN (A,IA,N,B,APVT,X)
C     DO 20 I=1,N
      DO 20 I=1,N
        DO 20 I=1,N
          IP = APVT(I)
          SUM = X(IP)
          X(IP) = X(I)
          IF (IW .EQ. 0) GO TO 15
          IM1 = I-1
          DO 10 J=IW,IM1
            SUM = SUM-A(I,J)*X(J)
          10 CONTINUE
          GO TO 20
          15 IF (SUM .NE. 0.0D0) IW = I
          20 X(I) = SUM
            SOLVE UX = Y FOR X
            DO 30 IB=1,N
              I = N+1-IB
              IP1 = I+1
              SUM = X(I)
              IF (IP1 .GT. N) GO TO 30
              DO 25 J=IP1,N
                SUM = SUM-A(I,J)*X(J)
              25 CONTINUE
              30 X(I) = SUM/A(I,I)
            RETURN
          END

          SUBROUTINE UERTST (IER,NAME)
          SPECIFICATIONS FOR ARGUMENTS
          INTEGER IER
          INTEGER NAME(1)
          SPECIFICATIONS FOR LOCAL VARIABLES
          INTEGER NAMESET(6),NAMUPK(6),NIN,NOUT,
          NAMEQ(6*1H) /
          LEVEL/4/,IEQDF/0/,IEQ/1H=/
          * DATA
          DATA
          DATA
          CALL USPKD (NAME,6,NAMUPK,NMTB)      GET OUTPUT UNIT NUMBER
          CALL UGETIO(1,NIN,IOUNIT)             CHECK IER
          END

          SUBROUTINE UGETIO(IOPT,NIN,NOUT)
          SPECIFICATIONS FOR ARGUMENTS
          INTEGER IOPT,NIN,NOUT
          C   SPECIFICATIONS FOR LOCAL VARIABLES
          INTEGER NIND,NOUT
          DATA
          DATA
          CALL FIRST EXECUTABLE STATEMENT
          C   FIRST EXECUTABLE STATEMENT
          IF (IOPT.EQ.3) GO TO 10
          IF (IOPT.EQ.2) GO TO 5
          IF (IOPT.NE.1) GO TO 9005
          NIN = NIND
        END
      END
    END
  END
END
C   IF (IER.GT.999) GO TO 25
  IF (IER.LT.-32) GO TO 55
  IF (IER.LE.128) GO TO 5
  IF (LEVEL.LT.1) GO TO 30
    IF (IEQDF.EQ.1) WRITE(10UNIT,35) IER,NAMEQ,IEQ,NAMUPK
    IF (IEQDF.EQ.0) WRITE(10UNIT,35) IER,NAMUPK
    GO TO 30
  5 IF (IER.LE.64) GO TO 10
  IF (LEVEL.LT.2) GO TO 30
    IF (IEQDF.EQ.1) WRITE(10UNIT,40) IER,NAMEQ,IEQ,NAMUPK
    IF (IEQDF.EQ.0) WRITE(10UNIT,40) IER,NAMUPK
    GO TO 30
  10 IF (IER.LE.32) GO TO 15
    IF (LEVEL.LT.3) GO TO 30
    IF (IEQDF.EQ.1) WRITE(10UNIT,45) IER,NAMEQ,IEQ,NAMUPK
    IF (IEQDF.EQ.0) WRITE(10UNIT,45) IER,NAMUPK
    GO TO 30
  15 CONTINUE
    IF (LEVEL.LT.4) GO TO 30
    IF (IEQDF.EQ.1) WRITE(10UNIT,50) IER,NAMEQ,IEQ,NAMUPK
    IF (IEQDF.EQ.0) WRITE(10UNIT,50) IER,NAMUPK
    30 IEQDF = 0
    RETURN
  25 CONTINUE
    IF (LEVEL.LT.4) GO TO 30
    IF (IEQDF.EQ.1) WRITE(10UNIT,50) IER,NAMEQ,IEQ,NAMUPK
    IF (IEQDF.EQ.0) WRITE(10UNIT,50) IER,NAMUPK
    30 IEQDF = 0
    RETURN
  35 FORMAT(19H *** TERMINAL ERROR,10X,7H(IER = ,I3,
     1           20H) FROM IMSL ROUTINE ,6A1,A1,6A1)
  40 FORMAT(27H *** WARNING WITH FIX ERROR,2X,7H(IER = ,I3,
     1           20H) FROM IMSL ROUTINE ,6A1,A1,6A1)
  45 FORMAT(18H *** WARNING ERROR,11X,7H(IER = ,I3,
     1           20H) FROM IMSL ROUTINE ,6A1,A1,6A1)
  50 FORMAT(20H *** UNDEFINED ERROR,9X,7H(IER = ,I5,
     1           20H) FROM IMSL ROUTINE ,6A1,A1,6A1)
  C   SAVE P FOR P = R CASE
  C   P IS THE PAGE NAMUPK
  C   R IS THE ROUTINE NAMUPK
  C   55 IEQDF = 1
  DO 60 I=1,6
  60 NAMEQ(I) = NAMUPK(I)
  65 RETURN
END
C   SUBROUTINE UGETIO(IOPT,NIN,NOUT)
          SPECIFICATIONS FOR ARGUMENTS
          INTEGER IOPT,NIN,NOUT
          C   SPECIFICATIONS FOR LOCAL VARIABLES
          INTEGER NIND,NOUT
          DATA
          DATA
          CALL FIRST EXECUTABLE STATEMENT
          C   FIRST EXECUTABLE STATEMENT
          IF (IOPT.EQ.3) GO TO 10
          IF (IOPT.EQ.2) GO TO 5
          IF (IOPT.NE.1) GO TO 9005
          NIN = NIND
        END
      END
    END
  END
END

```

program

16

```
NOUTD = NOUTD  
GO TO 9005  
5 NIND = NIN  
GO TO 9005  
10 NOUTD = NOUT  
9005 RETURN  
END
```

Thu Jun 9 13:42:25 1994